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New Bid-Ask Spread Estimators from Daily High and Low Prices

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Abstract

In this paper, we introduce two low frequency bid-ask spread estimators using daily high and low transaction prices. The range of mid-prices is an increasing function of the sampling interval, while the bid-ask spread and the relationship between trading direction and the mid-price are not constrained by it and are therefore independent. Monte Carlo simulations and data analysis from the equity and foreign exchange markets demonstrate that these models significantly out-perform the most widely used low-frequency estimators, such as those proposed in [Corwin and Schultz \(2012\)](#) and most recently in [Abdi and Ranaldo \(2017\)](#). We illustrate how our models can be applied to deduce historical market liquidity in NYSE, UK, Hong Kong and the Thai stock markets. Our estimator can also effectively act as a gauge for market volatility and as a measure of liquidity risk in asset pricing.

Keywords: High-low spread estimator; effective spread; transaction cost; market liquidity

JEL Classification: C02, C13, C15

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1 Introduction

Estimating trading costs in the absence of recorded data is a problem that researchers of financial markets have been preoccupied with for some time. A bid-ask spread estimator that can be used on a range of market instruments with minimum input data and which is both accurate and efficient is the *sine qua non* in scholarly research when true trading cost data is unavailable. Scholars expend much effort in arriving at an estimator that resolves this issue of opaque trading cost data, and models have evolved considerably in the years since [Roll \(1984\)](#) originally formulated the concept. In this paper, we introduce two low frequency bid-ask spread estimators which can create estimates of these costs using the range between daily and two day high and low prices.

We also demonstrate how the models we propose significantly outperform those introduced by [Corwin and Schultz \(2012\)](#) (the CS estimator hereafter), [Abdi and Rinaldo \(2017\)](#) (the AR estimator hereafter) and the benchmark estimator introduced in [Roll \(1984\)](#). We perform tests using Monte Carlo simulations, and conduct empirical inquiry using foreign exchange and U.S. equity market data. Our models are designed along similar principles to the CS estimator in that it assumes that high and low prices are based on buy and sell transactions respectively. We present two models, the first, which we name the basic version uses a transaction range which is determined in part by the mid-price range and by the bid-ask spread. We posit that the mid-price range is a function of the time interval. Therefore, by comparing the ranges of transaction prices from the two different sampling frequencies, we can isolate the impact of the bid-ask spread. Our second model, which we refer to as the sophisticated version, which builds on that proposed in [Bleaney and Li \(2016\)](#) (the BL estimator hereafter). In the model we present, the bias that occurred as a result of feedback trading which is evident in the BL model has been used to link the one day and two day range in order to arrive at an estimation of the spread. We note that the bias that results through feedback trading is a function of the time interval. By comparing both the one and two day BL spreads we can obtain the estimates of the bid-ask spread. Our investigation of estimator performance is based upon the analysis of the mean and standard deviation of estimated errors alongside the correlation between these and the true spreads. In

addition we further the argument beyond simply relying on one single criterion such as correlation to indicate performance. Instead, we show that gauging performance on a range of indicators is the optimum path to choose for researchers who require an accurate measure of trading costs.

The estimation of accurate bid-ask spreads has for a long time been considered a significantly important part of market microstructure theory. Bid-ask spread estimators allow researchers and practitioners to develop trading strategies that incorporate an idea of the costs attached to each transaction. In turn, this allows the determination, with more accuracy, of the profitability that can be achieved from following specific trading rules. Understanding market liquidity is also important to researchers, so a precise estimation of the bid-ask spread offers a clearer picture of this market characteristic (Mancini et al. 2013; Banti et al. 2012). Excessive costs attached to accessing bid-ask spread data has meant that researchers increasingly rely on such estimators to aid their analysis of market activity. While some research indicates that the cheaper daily closing quoted bid-ask spread can be good proxy for the intraday spread (Holden and Jacobsen 2014, Chung and Zhang 2014 and Fong et al. 2017), unavailability of spread data is not simply a consequence of poor research budgets. Historical information on both quoted and true spreads is not always available. As a result, a strong performing estimator model can serve even well-resourced researchers. Another possible use for these models arises from the fact that bid-ask spreads can influence measures associated with price volatility so scholars analysing this metric can use estimator models to arrive at an accurate measure of this widely researched market phenomenon (e.g. Bandi and Russell 2006).

Bid-ask spread estimation models need to satisfy certain requirements before they become useful to researchers. Models must be accurate, efficient in terms of having a low standard deviation of estimates and they must have low requirements on the type of data needed for computation. With accuracy and efficiency, the signal to noise ratio is an important consideration; the spread (signal) is more difficult to estimate when it is considerably smaller than the mid-price volatility levels (noise). Assets with high levels of liquidity demonstrate typically smaller spreads; however with more infrequently traded instruments, the bid-ask spreads can be quite large. Longer sampling intervals tend to display higher mid-price volatility levels, therefore, the signal to noise

ratio is smaller in more infrequent samples. This leads to a poorer performance in the accuracy and efficiency of estimators that have to rely on low sampling frequencies, as pointed out by [Bleaney and Li \(2015\)](#). In addition to the need for models to be both accurate and efficient, other barriers to inquiry may inhibit a model's usefulness. Constraints on accessing data imposed by availability or cost mean that models with more modest data requirements are of greater use to researchers. For instance, [Roll \(1984\)](#) requires just the transaction price of assets in order to apply the estimator, whereas [Huang and Stoll \(1997\)](#) require both the transaction price and the trade direction. [Corwin and Schultz \(2012\)](#) require the high low price range. [Abdi and Ranaldo \(2017\)](#) require the closing price and the high-low price range.

In this paper, we analyse the performance of the estimators through conducting a series of tests using both randomly generated and real data, the latter is taken from both the foreign exchange and equity markets. In most cases, both of our estimators outperform all the others tested. In comparison, the CS estimator is unstable as it only works well for equities but not for the remaining tests. The AR estimator exhibits high correlations with the true spread, while displaying a tendency to remain lower than those we uncover and also does not perform well in terms of the root mean square error (RMSE). Testing using simulation experiments produces a signal to noise ratio over 125000 months of generated data which ranges from 0.005 to 0.387. This covers most cases which have occurred in actuality. Both of our proposed models outperform the others mentioned in terms of both efficiency and accuracy. We also move beyond time series testing to investigate the cross sectional performance on a generated sample of 75000 data months, again at this level we find that our models outperform the others tested.

In the existing literature, bid-ask spread estimators are tested using the price data taken from the equities markets. An additional benefit offered by our models is that these are suitable for use both in the equities and foreign exchange markets because they are independent of the market structure. In this paper we run tests using data taken from both financial markets. In both types of markets the empirical tests are conducted using spreads calculated from tick by tick data as a benchmark, we find that our estimators again perform better than other models available to researchers.

In presenting the significant contributions of our work, the rest of our paper is or-

ganised as follows. Section 2 discusses existing bid-ask spread estimators, while section 3 introduces our new models. Section 4 reports the performance of these against that of the Roll, CS and AR estimators. Section 5 illustrates some applications of our estimator using equity markets and section 6 concludes.

2 Relevant bid-ask spread estimators

Spread estimator models are generally classified into one of four categories, the Roll, the LOT, the Effective Tick and more recently the High-low estimator. Each approach provides alternative methods which are based on the return autocovariance, the interval fractions in trade prices, the frequency of zero returns and the specific interval determined price range.

Roll (1984) firstly proposes an estimator for the bid-ask spread. This model was popularly received and generated considerable interest, giving rise to attempts to refine it further by later scholars. Roll's premise was to use return autocovariance to estimate the spread; underpinning this approach was the assumption that prices followed a random walk. It was also assumed that the closing stock price equalled its true value plus or minus half of the effective spread. The estimated spread could then be calculated as twice the square root of minus one multiplied by the autocovariance of the sample of daily returns. Prior research points out some problems with this approach, for instance, one concern is that the estimators produce results which can often underestimate the spread (Harris 1990). To deal with this autocorrelated mid-price return bias^a, George et al. (1991) suggest modifying the original Roll estimator. Similarly, Choi et al. (1988) introduce adjustments to the model in an attempt to deal with the problem of auto correlated trade directions. Stoll (1989) attempts to address the bias by taking the impact of inventory control and asymmetric information costs into account. To reach a general solution to the problem, Huang and Stoll (1997) incorporate each of the estimators above in one general model. However, gathering the data required to run this particular version of the model becomes a difficult process as trade direction data is also required. Hasbrouck (2004, 2009) suggests that a more accurate spread can be achieved through employing Gibbs estimation. Unlike Roll's approach

^aThe bias arises as the assumption that returns are random is not satisfied.

the computational requirements to employ Hasbrouck's estimators are considerably more intensive. The problem of normality in Hasbrouck's is addressed in [Chen et al. \(2016\)](#) who propose a non-parametric method to estimate the spread based on the Roll model. A further development is found in [Abdi and Rinaldo \(2017\)](#) who incorporate the CS model into the Roll to derive a new estimator. The correlation between its estimates and the true spread is higher than the CS estimator, but the RMSE is not significantly better than the CS.

Another estimator which scholars use to deal with the problem of the spread opacity due to the unavailability of data is proposed by [Lesmond et al. \(1999\)](#). Otherwise known as the LOT model, an effective spread is arrived at by considering the fraction of returns which are different from zero. This model has not quite reached the popularity levels of the Roll model as comparatively it tends not to perform as well in empirical testing ([Corwin and Schultz 2012](#)). [Holden \(2009\)](#) and [Goyenko et al. \(2009\)](#) put forward a more sophisticated approach which estimates spreads using effective tick measures based on the phenomenon of price clustering, a term describing the tendency for trade prices to occur most frequently on rounder price increments. However, results produced following testing on an extensive FX market data sample by [Karnaukh et al. \(2015\)](#) show that the LOT and effective tick estimators display only a weak relationship with true spreads.

The high-low spread estimator introduced by [Corwin and Schultz \(2012\)](#) is a later addition to the toolkit of estimators. Despite being relatively new, it is used extensively in recent literature because testing shows that it satisfies the estimator requirements to a greater extent than previous innovations ([Corwin and Schultz 2012](#), [Holden and Jacobsen 2014](#) and [Karnaukh et al. 2015](#)). The model is derived using the high and low prices of an asset over two day and daily horizons. The relative ease with which input data can be accessed further enhances the appeal of the estimator. The estimator we introduce here operates in the spirit of the CS model by adopting similar assumptions that the high (low) prices are most likely buy (sell) orders.

3 The High-low estimators

The general structure of both of our estimators can be expressed in the following equation:

$$\frac{E \left(\sqrt{2} \cdot X_{daily} - X_{today} \right)}{\sqrt{2} - 1} \quad (1)$$

Where X is the price range or the spread estimated by the [Bleaney and Li \(2016\)](#) estimator which is based on the variance of mid-price returns. The innovations we introduce for both the basic and sophisticated models are discussed in the following subsections.

3.1 The basic high-low estimator (the BHL model)

When estimating a model using the high and low transaction prices, the first characteristic that we can note is that the range increases as the time interval widens because the mid-price volatility grows in proportion to this. The other factor contributing to the range is the bid-ask spread, however this is independent of the time interval. Therefore, it is possible to extract the bid-ask spread by calculating the difference between the high and low transaction prices, whilst considering inconsistencies that may arise as a result of volatility.

In order to accomplish this, we assume that the mid-price, denoted as M_t , follows a one-dimensional Wiener process. The link between the unobserved mid-price and the observed transaction price (s_t) is given through the following equation.

$$s_t = M_t + \frac{SP}{2} \cdot BS_t \quad (2)$$

Where BS_t is the trade indicator showing 1 (−1) for a buyer (seller) initiated trade. The relationship between the daily high mid-price (H_t^M) and the daily high transaction price (H_t^T) as well as the link between the daily low mid-price (L_t^M) and the daily low transaction price (L_t^T) are demonstrated in the following set of equations:

$$\begin{aligned} H_t^T &= H_t^M + \frac{SP}{2} \cdot BS_t & L_t^T &= L_t^M + \frac{SP}{2} \cdot BS_t \\ TH_t^T &= TH_t^M + \frac{SP}{2} \cdot BS_t & TL_t^T &= TL_t^M + \frac{SP}{2} \cdot BS_t \end{aligned} \quad (3)$$

Where T and M represent the transaction and mid-price respectively. TH and TL denote the high and low prices over a two day window.

We can eliminate the need to establish trade direction by assuming that the highest (lowest) prices are buy (sell) orders. [Bleaney and Li \(2015\)](#) support this assumption showing that in more than 99% of cases, the daily highest transaction price is a buyer initiated trade whereas its lowest corresponding value is seller initiated. Formally, it can be represented as:

$$BS_t = \begin{cases} 1 & \text{if } s_t = H_t^T \\ -1 & \text{if } s_t = L_t^T \end{cases} \quad (4)$$

The daily and two-day ranges of transactions represent the difference between the highest and lowest prices. Formally, these expectations are given as:

$$\begin{aligned} Range_{t,daily}^T &= H_t^T - L_t^T \\ &= \left(H_t^M + \frac{SP}{2} \cdot BS_t \right) - \left(L_t^M + \frac{SP}{2} \cdot BS_t \right) \\ &= \left(H_t^M + \frac{SP}{2} \right) - \left(L_t^M - \frac{SP}{2} \right) \\ &= (H_t^M - L_t^M) + SP \\ &= Range_{t,daily}^M + SP \end{aligned} \quad (5)$$

$$\begin{aligned} Range_{t,twoday}^T &= TH_t^T - TL_t^T \\ &= \left(TH_t^M + \frac{SP}{2} \cdot BS_t \right) - \left(TL_t^M + \frac{SP}{2} \cdot BS_t \right) \\ &= \left(TH_t^M + \frac{SP}{2} \right) - \left(TL_t^M - \frac{SP}{2} \right) \\ &= (TH_t^M - TL_t^M) + SP \\ &= Range_{t,twoday}^M + SP \end{aligned} \quad (6)$$

Where $Range_{t,daily}^T$ and $Range_{t,twoday}^T$ are daily and two-day ranges respectively. The equations above demonstrate our earlier suggestion that the range of transaction prices is influenced by volatility in both the mid-price and the bid-ask spread. Taking expectations of both sides, the equations become:

$$E \left(Range_{daily}^T \right) = E \left(Range_{daily}^M \right) + SP \quad (7)$$

$$E \left(Range_{twoday}^T \right) = E \left(Range_{twoday}^M \right) + SP \quad (8)$$

The left hand side of Equations (7) and (8) can be calculated from observed transaction prices. With the unobserved terms, the expected ranges of daily and two-day mid-prices can be eliminated, allowing us to extract the bid-ask spread. [Parkinson \(1980\)](#) shows that if the mid-price follows a one-dimensional Wiener process, its expected

range is an increasing function of the sampling time interval and its diffusion. A long sampling time interval or large diffusion will lead to a wider range. Formally, the expectation of the range of mid-prices can be calculated through the following equation:

$$E \left(Range^M \right) = \sqrt{\frac{8D \cdot ti}{\pi}} \quad (9)$$

Where D is the diffusion of mid-prices in a unit time interval (ti). If this period is one day, the expectations for daily and two-day ranges are given through the following equations:

$$E \left(Range_{daily}^M \right) = \sqrt{\frac{8D}{\pi}} \quad (10)$$

$$E \left(Range_{twoday}^M \right) = \sqrt{\frac{8D}{\pi}} \cdot \sqrt{2} \quad (11)$$

Therefore, the expectation is that the two-day range is $\sqrt{2}$ times that of the daily range. Formally, the relationship is expressed through the following equation:

$$E \left(Range_{twoday}^M \right) = \sqrt{2} \cdot E \left(Range_{daily}^M \right) \quad (12)$$

From Equations (7), (8) and (12), we can solve the bid-ask spread (SP), because we have three equations and three unknown variables. We solve Equation (8) through deducting $\sqrt{2}$ times each side of Equation (7):

$$\begin{aligned} & E \left(Range_{twoday}^T \right) - \sqrt{2} \cdot E \left(Range_{daily}^T \right) \\ & = E \left(Range_{twoday}^M \right) + SP - \sqrt{2} \cdot \left[E \left(Range_{daily}^M \right) + SP \right] \end{aligned} \quad (13)$$

When we substitute Equation (12) into Equation (13), and rearrange the yields, the estimate of the bid-ask spread becomes:

$$SP = \frac{E \left[\sqrt{2} \cdot \left(Range_{daily}^T \right) - \left(Range_{twoday}^T \right) \right]}{\left(\sqrt{2} - 1 \right)} \quad (14)$$

Equation (14) represents the basic estimator which we propose in this paper (BHL hereafter); this is an expectation of the linear function of the daily and two-day high and low transaction prices. One of its key features is that it is unbiased and easy to compute. It outperforms the CS estimator because it produces an unbiased result while remaining linear. Using BHL, it is possible to increase the number of observations in order to obtain a more accurate and efficient estimate of the spread. This is because

statistical errors and noise can be eliminated from large sample sizes; this is not the case for non-linear estimators (Bleaney and Li 2015). Furthermore, the estimates remain consistent across a variety of sampling periods. This suggests that when higher sampling frequency data becomes available we can use this to obtain a more accurate and efficient estimate because, as Bleaney and Li (2015) suggest, the noise (the price volatility) is relatively low in comparison with the bid-ask spread.

Similar to the CS estimator, we can also estimate the daily diffusion, which is expressed as D , using the same process. This is represented in the following equation:

$$E \left(Range_{daily}^M \right) = \frac{E \left(Range_{twoday}^T \right) - E \left(Range_{daily}^T \right)}{(\sqrt{2}-1)} = \sqrt{\frac{8D}{\pi}}$$

$$D = \frac{\pi}{8} \left[\frac{E \left(Range_{twoday}^T \right) - E \left(Range_{daily}^T \right)}{(\sqrt{2}-1)} \right]^2 \quad (15)$$

3.2 The sophisticated high-low estimator

Our sophisticated estimator (the SHL model hereafter) introduces innovations to the design proposed by Bleaney and Li (2016). The BL model shows that the corrected estimated spread allows for the maximisation of the covariance of the mid-price. The BL estimator is distinctive in that it outperforms Roll (1984), Huang and Stoll (1997), Corwin and Schultz (2012) and Hasbrouck (2009) estimators, following extensive testing. The disadvantage with the BL model is that its computation requires both the transaction prices and information on the direction of trade. However it is often the case that researchers don't have access to this information.

The SHL model differs from the BL estimator in two respects. First, SHL uses estimated spread which maximises the variance of mid-prices; second, SHL introduces the assumption that the highest prices recorded daily are ask-prices and the lowest are bid-prices. Through these assumptions we can lower the data requirements for the model and allow the estimator to operate using only the highest and lowest transaction prices in the estimation window.

Following Bleaney and Li (2016), we let A be a set of all conjectures of the true spread where the symbol \sim represents conjectural values.

$$A = \left\{ \widetilde{SP}_1, \widetilde{SP}_2, \dots, \widetilde{SP}_n \right\} \quad (16)$$

B denotes a set of variances of mid-price returns, the conjecture being that the true spread is taken to be:

$$B = \{Var_1, Var_2, \dots, Var_n\} \quad (17)$$

Where

$$Var_i = Var \left[\Delta \tilde{M}(\tilde{SP}_i)_t \right] \quad (18)$$

Based on these settings we can propose the following:

Proposition 3.1 *When the components of the spread do not include feedback trading, inventory control or asymmetric information, we can consider that the spread and its estimates, and thus the estimated errors, are either serially independent or fixed. If an estimate of the spread $\tilde{SP}_i \in A$ corresponds to $Var_i = \max(B)$, it equals the true spread which is then denoted as: $\tilde{SP}_i = SP$.*

Proof The full proof is in the appendix. The variance of two adjacent conjectures of mid-price returns is:

$$\begin{aligned} Var_i &= Var \left[\Delta \tilde{M}_t \right] \\ &= E \left\{ \left[\Delta \tilde{M}_t - E \left(\Delta \tilde{M}_t \right) \right]^2 \right\} \end{aligned} \quad (19)$$

We will assume that the expectation of the value of the conjectural mid-price is zero. Thus, the equation above can be rewritten as:

$$\begin{aligned} Var_i &= Var \left(\Delta \tilde{M}_t \right) \\ &= E \left(\Delta \tilde{M}_t^2 \right) \\ &= E \left[\left(\Delta M_t + \frac{1}{2} \Omega BS_t - \frac{1}{2} \Omega BS_{t-1} \right)^2 \right] \end{aligned} \quad (20)$$

Where Ω denotes the conjectural error which represents the difference between the conjectural mid-price and the true mid-price, alternately expressed as the difference between the conjectural spread and its true value. Formally, Ω is given as:

$$\Omega = \Delta \tilde{M}_t - \Delta M_t = \tilde{SP}_i - SP \quad (21)$$

The assumptions of this proposition imply that BS is independent of ΔM at all observation points, therefore many of the terms in (20) such as $E(\Delta M_t BS_t)$ and $E(\Delta M_{t-1} BS_t)$

equate to zero. The variable BS is a binary variable (1 or -1), thus $E(BS_{t-1}^2) = 1$. Finally we obtain:

$$\begin{aligned}
Var_i &= Var(\Delta\tilde{M}_t) \\
&= E(\Delta\tilde{M}_t^2) \\
&= E\left[\left(\Delta M_t + \frac{1}{2}\Omega BS_t - \frac{1}{2}\Omega BS_{t-1}\right)\left(\Delta M_t + \frac{1}{2}\Omega BS_t - \frac{1}{2}\Omega BS_{t-1}\right)\right] \\
&= E\left(\Delta M_t^2 + \frac{1}{2}\Omega^2\right)
\end{aligned} \tag{22}$$

The final step of Equation (22) given above is the quadratic polynomial of the expectation of the error of the conjecture. For a given series, the first term $E(\Delta M_t^2)$ is a constant. We can surmise directly from this that when the error is zero (i.e. $\Omega = 0$), the second term $\frac{1}{2}\Omega^2$ is zero. Furthermore, when $\Omega = 0$, there is a global extreme for the right hand side polynomial in the final step, and symmetrically, the left hand side of the equation $Var_i = Var(\Delta\tilde{M}_t)$ is also at the extreme value. Formally this can be expressed as:

$$arg \max_{\Omega} Var(\Delta\tilde{M}_t) = 0 \tag{23}$$

When the conjectural error is zero, the conjectural spread becomes the true spread:

$$\tilde{SP}_i = SP + \Omega = SP \tag{24}$$

Therefore the conjectural spread which maximises the covariance equals the true spread.

$$arg \max_{\tilde{SP}_i \in A} Var(\Delta\tilde{M}_t) = SP \tag{25}$$

Q.E.D.

Figure 1 outlines the reasoning underpinning this proposition where for the purposes of economy we hold that the mid-price is fixed. The conjectural spread (\tilde{SP}_i) is less than the true spread. This allows us to estimate the conjectural mid-price \tilde{M} ; this is represented by the dotted line in Figure 1, and the true mid price and transaction price are both represented by unbroken lines. Also in Figure 1, A and B denote observed ask and bid prices, whereas M is the unobserved true mid-price.

At any one point we can only observe one price, either the bid or the ask. In Figure 1, three periods are displayed. In the period labelled $t - 2$, the bid price is recorded and in period labelled $t - 1$, the ask price is observed. In period $t - 2$, the conjectural

spread is lower than the true spread and the conjectural mid-price error is -0.5Ω , which is less than the true value. In period $t - 1$, the conjectural mid-price error is 0.5Ω , therefore this is greater than the true one. In the intervening period between $t - 2$ and $t - 1$, the direction of the trade shifts from sell to buy, and because of the conjectural error, we overestimate the mid-price return, formally we express this as:

$$\Delta\tilde{M}_{t-1} = \Delta M_{t-1} + \Omega = \Omega \quad (26)$$

In Figure 1, the hypothetical example shows that the variance of mid-price returns equates to zero because returns remain fixed. However the variance of conjectured mid-price returns is greater than zero. The reason for this is that in the case where the spread is underestimated, the conjectured mid-price fluctuates more than its true counterparts.

[Insert Figure 1 here]

According to the abovementioned proposition, we find that the true spread maximises the variance of conjectural mid-price returns and can be expressed as follows:

$$\begin{aligned} \text{Var}(\Delta\tilde{M}_t) &= E(\Delta\tilde{M}_t^2) \\ &= E\left[\left(\Delta s_t - \frac{\tilde{SP}}{2}\Delta BS_t\right)^2\right] \\ &= E(\Delta s_t^2) - \tilde{SP} \cdot E(\Delta s_t \Delta BS_t) + \frac{\tilde{SP}^2}{4} \cdot E(\Delta BS_t^2) \end{aligned} \quad (27)$$

Using first order conditioning, we find that the estimated spread satisfies the following equation:

$$-E(\Delta s_t \Delta BS_t) + \frac{1}{2}\tilde{SP} \cdot E(\Delta BS_t^2) = 0 \quad (28)$$

$$SP = \tilde{SP} = \frac{2E(\Delta s_t \Delta BS_t)}{E(\Delta BS_t^2)} \quad (29)$$

Equation (29) is now the variance version of the BL estimator, thereby reflecting one of the suggested innovations that we propose in this paper.

In order to allow Equation (29) to become operational, we must introduce the following processes. On each day, we pick either the high or low price randomly to create a trail series of prices (s_t) and use Equation (4) to determine the trade direction: a buy order when s_t is the high price and a sell order when s_t is the low price. We can then calculate the estimated spread using Equation (29). In the same manner as [Corwin](#)

and Schultz (2012), we calculate an estimate of spread using the two-day high and low prices SP_{twoday} .

However, Equation (4) creates the link between order flow and price when only high and low prices are used. When the covariance between order directions and mid-price returns is non-zero, the BL estimator is biased and the error is expressed as $E(BS_t \cdot \Delta M_t)$ ^b. Therefore, when high and low prices and relevant trade directions are used, the BL estimator significantly overestimates the spread. It is invariably the case that the estimated spread will contain errors, these nevertheless can be offset if we compare the estimates using daily and two-day data ($E(SP_{daily})$ and $E(SP_{twoday})$). This is because it is possible to predict the relationship between errors from daily and two-day estimates.

The true spread is then taken to be the estimated spread minus the error. Formally, the true spread is given as follows:

$$\begin{aligned} SP &= SP_{daily} - E(BS_{daily} \cdot \Delta M_{daily}) \\ SP &= SP_{twoday} - E(BS_{twoday} \cdot \Delta M_{twoday}) \end{aligned} \quad (30)$$

Where the subscripts "daily" and "twoday" represent the sampling frequencies. SP_{daily} and SP_{twoday} are the BL estimates using daily and two-day high low data respectively. Following the discussion in the previous section, the relationship between daily and two-day ranges can be used to eliminate the estimated error above. The errors are in fact half of expected ranges of daily and two-day ranges and are expressed as follows (proof can be found in the appendix):

$$\begin{aligned} E(BS_{daily} \cdot \Delta M_{daily}) &= \frac{1}{2}E(H_t - L_t) = \frac{1}{2}E(Range_{daily}^M) \\ E(BS_{twoday} \cdot \Delta M_{twoday}) &= \frac{1}{2}E(TH_t - TL_t) = \frac{1}{2}E(Range_{twoday}^M) \end{aligned} \quad (31)$$

Following steps similar to the process outlined in section 3.1, we substitute Equations (31) and (12) into Equation (30), the rearrangement yields the spread for one trial series. This trial estimate is given as:

$$SP_{onetrial} = \frac{\sqrt{2} SP_{daily} - SP_{twoday}}{\sqrt{2} - 1} \quad (32)$$

^bThe feedback trading bias is discussed in Bleaney and Li (2016). Proofs are given in the appendix of this paper.

We repeat the trail series creation and estimation process over 1000 times, the mean of trail estimates becomes our estimate of spread. Although this process is computationally intensive, this makes little practical difference given the power of the current stock of computers available to researchers.

$$SP = E \left(\frac{\sqrt{2} SP_{daily} - SP_{twoday}}{\sqrt{2} - 1} \right) \quad (33)$$

Equation (33) is the sophisticated high low estimator (SHL). Theoretically, SHL should produce more accurate results than its BL counterpart. Unlike the BL model, the SHL estimator will not be influenced by feedback trading and the estimates produced will be unbiased.

When the ratio of the spread to the standard deviation of mid-prices is small, some trail estimates in Equation (32) could be negative. In order to avoid a negative result in Equation (33) we let all negative trial estimates equal zero. However, when we do this, the estimates produced by SHL might overestimate the spread. The simulation experiments we conduct in the next section show that it will not be an issue when the ratio becomes larger.

We can also estimate the daily diffusion, which is expressed as D , from Equations (30) and (31) and using the same process. This is represented in the following equations:

$$\begin{aligned} SP_{twoday} - SP_{daily} &= E (BS_{twoday} \cdot \Delta M_{twoday}) - E (BS_{daily} \cdot \Delta M_{daily}) \\ &= \frac{1}{2} E (Range_{twoday}^M) - \frac{1}{2} E (Range_{daily}^M) \\ &= \frac{(\sqrt{2}-1)}{2} \sqrt{\frac{8D}{\pi}} \end{aligned} \quad (34)$$

The rearrangement of the equation above yields an expression of daily diffusion as follows:

$$D = \frac{\pi}{2} \cdot \left[E \left(\frac{SP_{twoday} - SP_{daily}}{\sqrt{2} - 1} \right) \right]^2 \quad (35)$$

4 Comparison of the estimators

In this section, we examine the performance of a range of estimators. Using empirical tests, we gauge how BHL, SHL, Roll, CS and AR perform in addition to a number of equally weighted combinative models. Currently, the range of estimators available to

researchers is wide, but we focus on these models for two reasons. The first is that we wish to contrast the performance of our proposed models (BHL and SHL) with that of the best performing estimator available, the CS model. [Corwin and Schultz \(2012\)](#) demonstrate that the CS model outperforms all other low frequency estimators in terms of accuracy and efficiency. [Holden and Jacobsen \(2014\)](#) and [Karnaugh et al. \(2015\)](#) also show similar results to the model originators. We also choose the Roll model as this has traditionally been the benchmark for estimator performance. Researchers less familiar with the relatively recent CS model can understand how our models perform in comparison. Finally, we select the model proposed in [Abdi and Ranaldo \(2017\)](#) because it is the latest development of the spread estimator and is directly related to both the Roll and CS models. Our motivation behind including the combinatory models relates to the tendency for some estimators to over(under)estimate the spread. Combination of estimators has been shown in ([Holden 2009](#))^c to perform better in terms of accuracy. The data we use to test each of the models is tick by tick equity prices and foreign exchange rates; these are sourced from TAQ and Hotspot and DataStream respectively. In addition to using real world data testing, simulation experiments were also carried out. Our findings show that in general our BHL and SHL estimators outperform all other estimators included in the study.

4.1 Comparison strategy

Following testing on each of the estimators, the results are reported using average relative estimated errors together with their root mean square and standard deviation values. Formally, the relative error is defined as follows:

$$Rel - Err = \frac{Estimates - Spread}{Spread} \quad (36)$$

The average relative error (Rel-Error-Mean) reports the mean difference between the estimated and true spread, indicating where possible bias may exist in the estimators. When Rel-Error-Mean is positive (negative) it suggests that the models over (under) estimate the spread. Good estimates are those with 'close-to-zero' relative error aver-

^cCombination models tested in [Holden \(2009\)](#) and used here for testing are explained in Table 1.

ages. Formally, this is presented as:

$$Rel - Err - Mean = E (Rel - Err) \quad (37)$$

The standard deviation of the relative estimated errors (Rel-Err-Std) is also reported and provides a measure for the efficiency of the estimates. Good estimates have low Rel-Err-Stds. Formally, this is expressed as:

$$Rel - Err - Std = Std.Dev (Rel - Err) \quad (38)$$

Finally, the RMSE is the most widely used criteria by which to judge the performance of the estimators. Therefore we follow this trend in analysing how our models perform. Formally, RMSE is given as:

$$RMSE = \sqrt{E [(Rel - Err)^2]} \quad (39)$$

4.2 Simulation experiments

In this section we report the results of a number of simulations designed to test the relative strength of each measure. We find that the estimators proposed in this paper outperform the other models in terms of accuracy and efficiency.

4.2.1 Estimation under various 'signal to noise' ratios

It is difficult for an estimator to isolate the bid-ask spread from transaction prices when the volatility of mid-prices is relatively large. We test the performance of the estimators under various 'signal to noise' values which are the ratios of the spread to the standard deviation of the mid-price. The 'signal to noise' ratio is low for heavily traded equities and major currency pairs because the liquidity levels are consistently high, when the assumption is that mid-prices and order directions are random. We allow the standard deviation of one-minute mid-price returns to be 0.005 (about 0.19 daily). We consider six bid-ask spreads ranging from 0.001 to 0.3. The 'signal to noise' ratio extends from 0.00527 to 1.58 on a daily basis. In comparison, [Corwin and Schultz \(2012\)](#) test their model using the ratios which begin at 0.167 and end at 3.33; therefore the performance hurdles we employ to evaluate our estimators are more difficult to overcome.

Our simulation experiments are therefore more challenging and mirror real market conditions. For example, assuming that there are 20 trading days in a month, we compare the estimates of 25000 months. Formally, the data generation system is given as:

$$\begin{aligned}
 s_t &= M_t + \frac{SP}{2} \cdot BS_t \\
 BS_t &\sim B(1, 0.5) \\
 \Delta M_t &\sim N(0, 0.05) \text{ (one - minute)} \\
 SP &= \begin{cases} 0.001 \\ 0.002 \\ 0.006 \\ 0.010 \\ 0.030 \end{cases}
 \end{aligned}$$

Tables 1 to 5 report, the testing using various versions of our BHL, SHL and CS models^d. In general, those estimates are more accurate and efficient from the top left to the bottom right, as the ratio (*True spread/Midstd*) increases (from 0.00527 to 0.387) as the number of observations increases^e; this is consistent with findings of [Bleaney and Li \(2015, 2016\)](#). By setting negative trials and results to be zero, the BH3, SHL2, CS3, AR and Roll estimators demonstrate significant bias. For example, the relative error of SHL2 is 74.97% and those of BH3, CS3 and AR are 330%, 234% and 155% respectively when the ratio is 0.158 (The left panel of Table 5). If the ratio is 0.387 (The right panel of Table 5), the relative error of SHL2 is -0.593% which is close to zero, and therefore demonstrates the power of the model. For the other estimators, the relative errors of SHL1, BHL1 and BHL2 are less than 5% when the ratio is greater than 0.0258. According to the second column of Table 3, the average ranking of all simulation experiments in section 4.2.1 suggests that the combination of SHL2 and BHL1 is the best performing estimator. In terms of the performance of single rather than combined estimators, SHL2 offers the best results and is the second best performer from the entire array of models.

[Insert Tables 1 to 5 here]

^dSee the caption of Table 1 for full details of the versions of the estimators used.

^eOutliers of relative errors, the highest and lowest 1% of the relative estimated errors, are trimmed off before further calculation. We also test the cases of full sample and the case where the trimming is at the 0.05% and 2% level, the results produced are similar.

4.2.2 Cross-sectional properties of the estimators

In this section, the cross-sectional properties of the estimators are examined. In contrast to the previous section, the bid-ask spreads are assumed to vary each month and are evenly distributed from 0.002 to 0.0177. We also break the full sample into five groups according to the mean of the bid-ask spread. Then, we can examine the cross-sectional performances of the estimators across the five ranges of spread. The other parameters in the data generation process are the same as in the previous section.

Formally, the data generation system is given as:

$$\begin{aligned}
 s_t &= M_t + \frac{SP}{2} \cdot BS_t \\
 BS_t &\sim B(1, 0.5) \\
 \Delta M_t &\sim N(0, 0.05) \text{ (one - minute)} \\
 SP &= \begin{cases} \text{from 0.001 to 0.00513} \\ \text{from 0.00513 to 0.00829} \\ \text{from 0.00829 to 0.0114} \\ \text{from 0.0114 to 0.0146} \\ \text{from 0.0146 to 0.0177} \end{cases} \quad (40)
 \end{aligned}$$

The results of the simulation experiments are reported in Tables 6 to 11. The correlations between the true and estimated spreads are also reported. Table 6 reports the pooled results while the other panels are represented in the equation above according to each grouping of spreads. In the case of spreads where the range is between 0.01 and 0.03, the correlations reported are quite weak.

In the pooled case, although CS3 has a slightly stronger correlation than SHL2 with values of 0.136 and 0.127 respectively, it reports a much higher value for RMSE at 12.32 than the SHL2 value which is 5.51. In terms of correlation, the best performers are CS3, SHL2 and BHL3. From the third column of Table 12, it is evident from the average ranking of all simulation experiments that the combination of SHL2 and BHL1 and that of SHL2 and CS2 are the best performing estimators. For single models, SHL2 shows the best performance and is placed third in rank overall.

Table 12 shows a summarised average ranking for all simulations. According to the first column, the average ranking of all cases of simulation experiments suggests that the combination of SHL2 and BHL1 is the best performing estimator. SHL2 is the best

performing single estimator and takes second place overall. The other combinations outperform the other single estimators. The remaining alternate versions of our new models (BHL1, BHL2, SHL1) perform better than all the versions of the CS and Roll estimators.

[Insert Tables 6 to 12 here]

4.3 Comparisons in foreign exchange markets

In this section, we use our chosen estimators to gauge historical spreads for the foreign exchange markets. To do this, we test the estimators using closing prices and quoted spread data of 22 currency pairs in a sample beginning in January 1990 and finishing in December 2016, using data extracted from DataStream. We also use the prices and effective spreads of 12 Currency pairs in a sample dating from December 2015 to August 2016, this data is taken from Hotspot. In testing on both currency samples our sophisticated high-low estimator outperforms all others employed in the test.

4.4 DataStream 22 currency pairs 1990-2016

In this subsection, we evaluate the performance of the estimators using daily high and low prices and closing bid-ask spread data of 22 currency pairs taken from DataStream. For brevity, we report only the pooled results and not the result of the individual tests on each currency pair. The times series properties of the estimators are reported in Table 14 through the average performance ranking of the models for each pair examined; of these, SHL2 and the combination of SHL2 and BHL1 are the best performers.

According to Table 13, the combination of SHL2 and BHL1 outperforms the other estimators in terms of RMSE. Although CS2 has the lowest average estimated error (17.05%), its standard deviation of 2.315 is large relative to the others tested. SHL2 has a lower standard deviation than the aforementioned combination but a greater error (90.49%) and thus it takes second place in terms of performance. CS3 and BHL3 report high correlations between true and estimated spreads with the values of 0.876 and 0.869 respectively. However, the errors and the standard deviations of CS3 and BHL3 are much bigger than the others, and therefore CS3 and BHL3 perform poorly according to RMSE. SHL2 performs quite well in comparison, with a cross-sectional correla-

tion figure of 0.726, but reports a much lower error and standard deviation figure than CS3. Table 14 also suggests that the combination of SHL2 and BHL1 is the best choice while SHL2 provides a good alternative. Table 15 reports currency-by-currency average cross-sectional correlations between true and estimated spreads generated by each of the estimators. All the estimators exhibit very high correlations. BHL3's correlation of 0.95 is the highest. CS3 and AR's correlation of 0.948 are the second strongest. SHL2 exhibits a moderate correlation of 0.864. Table 15 also reports currency-by-currency time series correlations between true and estimated spreads generated by each of the estimators. Similar to those reported in Corwin and Schultz (2012), all the estimators exhibit weaker correlations than in the cross-sectional case. BHL3, CS3 and SHL2 exhibit the highest correlations, which are 0.271, 0.265 and 0.198 respectively.

[Insert Tables 13 to 15 here]

Figure 2 illustrates an example of the estimates and the actual closing quoted spread in the form of the USD/JPY currency pair over a 50 month period. We can see that all estimators, except for SHL2, show negative estimates. CS2 and the combinations appear more volatile than SHL2, the combination of SHL2 and BHL1 has the lowest average estimated error.

[Insert figure 2 here]

4.4.1 Hotspot 12 currency pairs Dec 2015-Aug 2016

In this section, we evaluate the performance of the estimators using daily high and low prices and the effective time-weighted bid-ask spread data of 12 currency pairs sourced from Hotspot. Results are reported in Table 16. Hotspot is a large electronic communication network (ECN) platform for foreign exchange transactions conducted worldwide. We extract quotes and transaction data similar to that taken from the TAQ database. Trade volume weighted effective spreads are calculated for each pair over time the sample period begins in December 2015 and ends in August 2016. Spreads are arrived at through the matching of quote and transaction data. The trade volume weighted effective spread can be formally expressed as:

$$\begin{aligned} & 2 \cdot (s_t - M_{t-1}) \quad \text{for buyer initiated trades} \\ & 2 \cdot (M_{t-1} - s_t) \quad \text{for seller initiated trades} \end{aligned} \tag{41}$$

In order to reduce the possibility of errors in the data we eliminate outliers^f and negative effective spreads. Table 16 displays the results of the pooled case where the 12 currencies over the entire sample period are examined. SHL2 is the best performing model in terms of RMSE, although its estimated error is very high (754%) but is similar to the others models. The standard deviation of SHL2 is the lowest (7.66) among the estimators tested, where each model either significantly over or under-estimates the spread. The currency-month pooled correlation coefficients for the estimated and true spreads for AR and BHL3 are 0.73 and 0.72 respectively; these are higher than the other models tested. However, their relative errors and standard deviations are greater in magnitude than the others. Therefore, AR and BHL3 are not the best performing estimators as the RMSEs place these at 13th and 14th in order of performance. The Roll estimator exhibits a high correlation coefficient (0.60), however, at least in half of the total cases, it obtains a negative number in the square root and we thus set it to zero. Although the combination of SHL2 and BHL1 is the second best in terms of RMSE, its correlation with the true spread is negative. SHL2's correlation coefficient is 0.24; this is acceptable in comparison with others. Also, its RMSE is the lowest amongst the models; therefore we rank this as the best performer. In Table 17, the average cross-sectional correlations between true and estimated spreads are reported across all currency pairs. AR and BHL3 exhibit the highest correlations, which are 0.8 and 0.77. SHL2's correlation is 0.636, performing slightly less well than AR and BHL3 in this instance but still at an adequate level. The average time series correlations between true and estimated spreads are reported across all currency pairs. BHL3 and AR exhibit the highest correlations, which are 0.51 and 0.48 respectively while SHL2's correlation is 0.04. The average time series correlations are much lower than those generated through cross-sectional analysis; this may be as a result of the time series being relatively short with its length being 12 months.

[Insert Tables 16 to 17]

^fOutliers are deemed to be those spreads which exceed the daily average by over 50 times.

4.4.2 Comparisons in equity markets TAQ data 2014

In this section, we use our chosen estimators to gauge spreads for the U.S equity market using the constituents of the S&P 1500 index as a sample. A snapshot of TAQ data, offers tick by tick pricing in 2014, which is used to calculate time-weighted quoted bid-ask spread and daily high and low prices. Tables 18 to 20 report the results of S&P 1500 (pooled case), S&P 600 (small cap), S&P 400 (mid cap), S&P 500 (large cap) stocks. In the pooled case, the SHL2 is the best performer in terms of RMSE results. It must be noted that all estimators significantly over or underestimate the spread. SHL2 displays the smallest estimated error but underestimates the spread by 59% on average, while the next best performing model (CS3) has an error of 104%. In contrast to a relatively poor performance with FX simulations, CS3 is ranked second out of all the estimators. The combination of SHL2 and BHL1 takes the third place. The pooled equity-month correlation coefficients of BHL3 and SHL2 are 0.78 and 0.74 respectively; this is significantly higher than the others and suggests a high correlation with the true spread. CS3 slightly outperforms SHL2 in the case of equities listed on the S&P600 because of its -1% average error. However, it represents an isolated example because the error of the CS3 estimator for simulations, equity and FX sample tests throughout the paper is high. Table 22 reports the average ranking of the small, mid and large cap equity group cases. It is apparent that SHL2 is still the best choice for estimator while CS2 and the combination of SHL2 and BHL1 produce results that could offer a good alternative. Table 23 reports the equity-by-equity cross-sectional correlations for each of the estimators. BHL3 and SHL2's correlations are 0.82 and 0.75 respectively; these are significantly higher than the others reported through the testing. AR's correlation of 0.70 is the third strongest. Table 23 reports average time series correlations of the estimators. BHL3 and CS3's correlations are 0.34 and 0.32. SHL2's correlation is 0.09.

[Insert Tables 18 to 23 here]

5 Application of SHL2

Moving beyond simulation, in this section, we demonstrate the application of SHL2 by using this estimator to gauge monthly average spreads for developed and emerg-

ing market stock exchanges. We find that the SHL2 acts as a good proxy for market liquidity as predictions of periods of intense uncertainty are often accompanied with low liquidity (high bid-ask spreads) levels in financial markets. In the samples we investigate, the data for the true spread is unavailable.

5.1 NYSE 1926-2015

Figure 3 shows the monthly average estimated spread of New York Stock Exchange (NYSE) from 1926 to 2015; this is generated by SHL2 using daily CRSP data. The spread was relatively large in the years before 1935. A further period of low liquidity can be observed from 1970 to 1992. As a result of the terrorist attack on 11 September 2001, the spread surged suddenly to over 5% in September 2001. In 2008 at the nadir of the global financial crisis, the SHL2 estimator recorded considerable lower liquidity levels through increased spreads, than in the years surrounding the event.

5.2 Non-US equity markets applications

Figures 4 to 6 show the monthly average estimated spreads for equities listed on the London, Hong Kong and Thai stock exchanges respectively. Data was obtained from Bloomberg.

We observe several increases in bid-ask spreads estimated by SHL2, i.e. transaction costs, around notable market events. For example, the average spread jumped to over 0.8% when the sterling crisis occurred in September 1992 (Figure 4). When the Asian financial crisis began in July 1997, transaction costs rose significantly in both the Thai and Hong Kong equity markets (Figures 5 and 6). The collapse of Lehman Brothers in September 2008 and the financial crisis which heralded drove a jump in spreads in almost all equity markets used in our samples. After the results of the Brexit referendum became clear in June 2016, transaction costs in the UK equity market also appeared to rise sharply.

[Insert figures 3 to 6 here]

The applications in this section suggest that SHL2 can act as an estimator that is sensitive enough to capture notable market events affecting transaction costs and as a consequence, liquidity levels. SHL2, as a spread estimator, can also be used as a

liquidity measure in asset pricing models in a similar manner to that demonstrated in [Corwin and Schultz \(2012\)](#) and [Abdi and Ranaldo \(2017\)](#).

6 Conclusion

In this paper, we introduce two new low frequency bid-ask spread estimators which estimate the bid-ask spread using daily and two-day high and low prices. We show that using similar input data, our estimators, in particular, the sophisticated version, significantly outperforms both the latest and the popular models such as [Abdi and Ranaldo \(2017\)](#), [Corwin and Schultz \(2012\)](#) and [Roll \(1984\)](#) in terms of accuracy, efficiency, as well as cross-sectional and time series correlations.

We test the performance of estimators using comprehensive Monte Carlo simulation experiments under various signal to noise ratios and different sampling frequencies. In addition, the cross-sectional properties of the estimators are also examined. Our estimators, BHL and SHL, appear to be unbiased throughout all tests carried out. By setting negative trials to zero which we label SHL2, we can obtain more efficient estimates; these are exhibited through lower standard errors. The results of simulation experiments suggest that our estimators outperform the AR, CS and Roll models. We demonstrate that SHL2 is the best single estimator in terms of accuracy and efficiency. We go further and test the performance of combinations of estimators against our own models and find that the AR, CS and Roll models also fail to match ours in performance. The combinations of the estimators are useful as using these can address the problems associated with errors which often appear for individual models. The combination of basic and sophisticated high-low estimators (BHL1 and SHL2) perform well and offer a good alternative to using the single estimators, thereby avoiding the associated error risk.

We then move beyond simulation experiments to study the models using real world data for both foreign exchange and equity markets. By comparing the closing bid-ask spread of 22 currency pairs over 26 years, we find that our SHL2 model outperforms all the others including the AR, CS and Roll models in terms of the root mean square error (RMSE). We arrive at the same conclusion when we ran tests using trade and quote data of 12 currency pairs over 9 months and for equities listed on the S&P 1500

throughout 2014. In terms of correlation, BHL3, AR, CS3 and SHL2 all performed well as estimators.

In general, our BHL and SHL are the best spread estimators. Researchers can choose the estimator according to their needs: BHL3 is good for cases where the high correlation is the only requirement and SHL2 can be used for other cases especially when accuracy and efficiency is of particular importance.

In order to demonstrate the effectiveness of our model (SHL2) through applications, we provide an illustration of how this can be applied. We generate the average monthly bid-ask spreads for the US, UK, HK and Thai equity markets. We show how the estimated spreads follow a pattern that is in line with our expectations in that the transaction costs increase sharply during crises periods.

Similar to the CS model, our estimators also obtain the estimates of daily mid-price diffusion at the same time as when the spreads are estimated. Because the spread and the diffusion are estimated together, a good spread estimator is also a good diffusion estimator. Thus, our sophisticated version model (SHL2) also offers the best diffusion estimates. As our estimators are not designed for a particular market structure, further research could test and apply our suggested estimators to the bonds, futures and option markets. In particular, these may be interesting for the over-the-counter markets where quote data can be difficult to obtain.

7 Appendix

7.1 Proof of Proposition 3.1

When the components of the spread do not include feedback trading, inventory control or asymmetric information, we can consider that the spread and its estimates, and thus the estimated errors, are either serially independent or fixed. If an estimate of the spread $\widetilde{SP}_i \in A$ corresponds to $Var_i = \max(B)$, it equals the true spread i.e. $\widetilde{SP}_i = SP$.

Proof The variance of two adjacent conjectures of mid-price returns is:

$$\begin{aligned} Var_i &= Var \left[\Delta \widetilde{M}_t \right] \\ &= E \left\{ \left[\Delta \widetilde{M}_t - E \left(\Delta \widetilde{M}_t \right) \right]^2 \right\} \end{aligned} \quad (42)$$

We will assume that the expectation of the value of the conjectural mid-prices is zero.

Thus, the equation above can be rewritten as:

$$\begin{aligned} Var_i &= Var \left(\Delta \widetilde{M}_t \right) \\ &= E \left(\Delta \widetilde{M}_t^2 \right) \\ &= E \left[\left(\Delta M_t + \frac{1}{2} \Omega BS_t - \frac{1}{2} \Omega BS_{t-1} \right)^2 \right] \\ &= E \left[\left(\Delta M_t + \frac{1}{2} \Omega BS_t - \frac{1}{2} \Omega BS_{t-1} \right) \left(\Delta M_t + \frac{1}{2} \Omega BS_t - \frac{1}{2} \Omega BS_{t-1} \right) \right] \\ &= E \left(\Delta M_t^2 + \frac{1}{2} \Omega BS_t \Delta M_t - \frac{1}{2} \Omega BS_{t-1} \Delta M_t \right) \\ &\quad + E \left[\frac{1}{2} \Omega BS_t \Delta M_t + \frac{1}{4} (\Omega BS_t)^2 - \frac{1}{4} \Omega^2 BS_t BS_{t-1} \right] \\ &\quad - E \left[\frac{1}{2} \Delta M_t \Omega BS_{t-1} + \frac{1}{4} \Omega^2 BS_t BS_{t-1} - \frac{1}{4} (\Omega BS_{t-1})^2 \right] \end{aligned} \quad (43)$$

Where Ω denotes the conjectural error which represents the difference between the conjectural mid-price and the true mid-price, alternately expressed as the difference between the conjectural spread and its true value. Formally, Ω is given as:

$$\Omega = \Delta \widetilde{M}_t - \Delta M_t = \widetilde{SP}_i - SP \quad (44)$$

The assumptions of this proposition imply that BS is independent of ΔM at all observation points, therefore many of the terms in (43) such as $E(\Delta M_t BS_t)$ and $E(\Delta M_{t-1} BS_t)$ equate to zero. Formally, we have:

$$\begin{aligned}
E(\Delta M_t BS_t) &= 0 \\
E(\Delta M_{t-1} BS_t) &= 0 \\
E(\Delta M_t BS_{t-1}) &= 0 \\
E(BS_t BS_{t-1}) &= 0
\end{aligned} \tag{45}$$

Furthermore, the variable BS is a binary variable (1 or -1), thus:

$$E(BS_{t-1}^2) = 1 \tag{46}$$

Finally we obtain:

$$\begin{aligned}
Var_i &= Var(\Delta \tilde{M}_t) \\
&= E\left(\Delta M_t^2 + \frac{1}{2}\Omega^2\right)
\end{aligned} \tag{47}$$

The final step of Equation (47) given above is the quadratic polynomial of the expectation of the error of the conjecture. For a given series, the first term $E(\Delta M_t^2)$ is a constant. We can surmise directly from this that when the error is zero (i.e. $\Omega = 0$), the second term $\frac{1}{2}\Omega^2$ is zero. Furthermore, when $\Omega = 0$, there is a global extreme for the right hand side polynomial in the final step, symmetrically, the left hand side of the equation $Var_i = Var(\Delta \tilde{M}_t)$ is also at the extreme value. Formally this can be expressed as:

$$arg \max_{\Omega} Var(\Delta \tilde{M}_t) = 0 \tag{48}$$

When the conjectural error is zero, the conjectural spread becomes the true spread:

$$\widetilde{SP}_i = SP + \Omega = SP \tag{49}$$

Therefore the conjectural spread which maximises the covariance equals the true spread.

$$arg \max_{\widetilde{SP}_i \in A} Var(\Delta \tilde{M}_t) = SP \tag{50}$$

Q.E.D.

7.2 Proof of feedback bias

When feedback trading exists, we have:

$$E(\Delta M_t BS_t) \neq 0 \tag{51}$$

Substituting Equations (45), (46) and (51) and into Equation (43), we can obtain:

$$\begin{aligned}
Var_i &= Var \left[\Delta \tilde{M}_t \right] \\
&= E \left[(\Delta M_t)^2 + \frac{1}{2} \Omega BS_t \Delta M_t - \frac{1}{2} \Omega BS_{t-1} \Delta M_t \right] \\
&+ E \left[\frac{1}{2} \Omega BS_t \Delta M_t + \frac{1}{4} (\Omega BS_t)^2 - \frac{1}{4} \Omega^2 BS_t BS_{t-1} \right] \\
&- E \left[\Delta M_t \frac{1}{2} \Omega BS_{t-1} + \frac{1}{4} \Omega^2 BS_t BS_{t-1} - \frac{1}{4} (\Omega BS_{t-1})^2 \right] \\
&= E \left(\Delta M_t^2 + \frac{1}{2} \Omega BS_t \Delta M_t \right) \\
&+ E \left(\frac{1}{2} \Omega BS_t \Delta M_t + \frac{1}{2} \Omega^2 \right) \\
&= E \left(\Delta M_t^2 + \Omega BS_t \Delta M_t + \frac{1}{2} \Omega^2 \right)
\end{aligned} \tag{52}$$

Substituting $\Omega = \widetilde{SP} - SP = 0$ into the equation above, we have:

$$\begin{aligned}
Var_i &= Var \left[\Delta \tilde{M}_t \right] \\
&= E \left[(\Delta M_t)^2 + \Omega BS_t \Delta M_t + \frac{1}{2} \Omega^2 \right] \\
&= E \left[(\Delta M_t)^2 + \left(\widetilde{SP} - SP \right) BS_t \Delta M_t + \frac{1}{2} \left(\widetilde{SP} - SP \right)^2 \right]
\end{aligned} \tag{53}$$

Using first order conditioning of Equation (53), we obtain:

$$SP = \widetilde{SP} - E (BS_t \Delta M_t) \tag{54}$$

Equation above suggests that when there is feedback trading, variance version of the BL estimator overestimates the spread.

7.3 Proof of Equation (30)

For each day, we choose at random either the daily high or low prices to calculate the daily price change. Thus, the probability of picking daily high (or low) price is 50% and there are four cases for the daily price changes with an equal likelihood which are as follows.

$$\Delta M_{daily} = \begin{cases} H_t - H_{t-1} & \text{with } \frac{1}{4} \text{ chance} \\ H_t - L_{t-1} & \text{with } \frac{1}{4} \text{ chance} \\ L_t - H_{t-1} & \text{with } \frac{1}{4} \text{ chance} \\ L_t - L_{t-1} & \text{with } \frac{1}{4} \text{ chance} \end{cases} \tag{55}$$

Thus $BS_{daily} \cdot \Delta M_{daily}$ is given as follows:

$$BS_{daily} \cdot \Delta M_{daily} = \begin{cases} BS_{daily} \cdot (H_t - H_{t-1}) & \text{with } \frac{1}{4} \text{ chance} \\ BS_{daily} \cdot (H_t - L_{t-1}) & \text{with } \frac{1}{4} \text{ chance} \\ BS_{daily} \cdot (L_t - H_{t-1}) & \text{with } \frac{1}{4} \text{ chance} \\ BS_{daily} \cdot (L_t - L_{t-1}) & \text{with } \frac{1}{4} \text{ chance} \end{cases} \quad (56)$$

When daily high (or low) price is picked, trading direction is known (Equation 4).

Formally, we have:

$$BS_{daily} \cdot \Delta M_{daily} = \begin{cases} [1 \cdot (H_t - H_{t-1})] & \text{with } \frac{1}{4} \text{ chance} \\ [1 \cdot (H_t - L_{t-1})] & \text{with } \frac{1}{4} \text{ chance} \\ [-1 \cdot (L_t - H_{t-1})] & \text{with } \frac{1}{4} \text{ chance} \\ [-1 \cdot (L_t - L_{t-1})] & \text{with } \frac{1}{4} \text{ chance} \end{cases} \quad (57)$$

Taking the expectation of $BS_{daily} \cdot \Delta M_{daily}$, we obtain:

$$\begin{aligned} E(BS_{daily} \cdot \Delta M_{daily}) &= \frac{1}{4} \cdot E(H_t - H_{t-1}) + \frac{1}{4} \cdot E(H_t - L_{t-1}) \\ &\quad - \frac{1}{4} \cdot E(L_t - H_{t-1}) - \frac{1}{4} \cdot E(L_t - L_{t-1}) \\ &= \frac{1}{2} E(H_t - L_t) \end{aligned} \quad (58)$$

7.4 A brief introduction to the AR, Roll and CS estimators

Researchers generally opt to use the Roll estimator and models⁸ derived from it because they are easy to program. The Roll estimator is given by the following equation.

$$SP = 2\sqrt{-cov(\Delta s_t, \Delta s_{t-1})} \quad (59)$$

According to [Corwin and Schultz \(2012\)](#), the CS estimator appears to be the best of low-frequency estimators including the [Lesmond et al. \(1999\)](#) estimator. Furthermore, our proposed model in this paper shares the same intuition with the CS estimator, therefore, the CS estimator is picked to examine. Squaring both sides of Equation (7), we have,

$$\begin{aligned} \left(Range_{daily}^T \right)^2 &= \left(Range_{daily}^M + SP \right)^2 \\ &= \left(Range_{daily}^M \right)^2 + 2 Range_{daily}^M \cdot SP + (SP)^2 \end{aligned} \quad (60)$$

⁸Related models include [Glosten and Harris \(1988\)](#), [Choi et al. \(1988\)](#), [Stoll \(1989\)](#), [George et al. \(1991\)](#), [Huang and Stoll \(1997\)](#), [Hasbrouck \(2004, 2009\)](#) and [Chen et al. \(2016\)](#)

Similarly, squaring both sides of Equation (8), we have:

$$\begin{aligned} \left(Range_{t\text{woday}}^T \right)^2 &= \left(Range_{t\text{woday}}^M + SP \right)^2 \\ &= \left(Range_{t\text{woday}}^M \right)^2 + 2 Range_{t\text{woday}}^M \cdot SP + (SP)^2 \end{aligned} \quad (61)$$

Corwin and Schultz (2012) assume that

$$\begin{aligned} E \left(Range_{t\text{woday}}^T \right)^2 &\approx E \left[\left(Range_{t\text{woday}}^T \right)^2 \right] \\ E \left(Range_{daily}^T \right)^2 &\approx E \left[\left(Range_{daily}^T \right)^2 \right] \end{aligned} \quad (62)$$

One could solve the spread from the equation system and obtains:

$$SP = \frac{2(e^\alpha - 1)}{1 + e^\alpha} \quad (63)$$

where

$$\alpha = \frac{\sqrt{2\beta} - \sqrt{\beta}}{3 - 2\sqrt{2}} - \sqrt{\frac{\gamma}{3 - 2\sqrt{2}}} \quad (64)$$

$$\beta = E \left\{ \sum_{J=0}^1 \left(Range_{daily,t+J}^T \right)^2 \right\}; \quad \gamma = \left(Range_{t\text{woday}}^T \right)^2 \quad (65)$$

When the spread is small, $SP \approx \alpha$. We may use Equation (64) to estimate the spread.

[Abdi and Ranaldo \(2017\)](#) model incorporates the CS model into the Roll estimator. Formally, it can be expressed as follows:

$$SP = 2\sqrt{(s_t - \eta_t)(s_t - \eta_{t+1})} \quad (66)$$

where η is the mid-point of the high and low prices. Formally, it is given by:

$$\eta_t = \frac{H_t + L_t}{2} \quad (67)$$

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Table 1: Simulation experiments: Comparison of the estimates over 25000 months

| | Daily (MidStd= 189.7*0.001) 20 observations per month | | | | | Four hours (MidStd =77.5*0.001) 120 observations per month | | | | |
|-------------------------------|--|--------------|-------------|---------|---------|---|--------------|-------------|--------|---------|
| | True spread= 1 (*0.001) | | | | | True spread/Midstd=0.0129 | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking |
| SHL1 | 0.953 | 17.76% | 41.177 | 41.176 | 8 | 0.899 | -8.763% | 6.513 | 6.514 | 7 |
| SHL2* | 35.348 | 3391% | 20.215 | 39.485 | 3 | 4.512 | 341% | 3.609 | 4.968 | 1 |
| (negatives to be zero) | | | | | | | | | | |
| BHL1* | 0.756 | -6.63% | 40.958 | 40.958 | 6 | 0.906 | -7.918% | 6.495 | 6.496 | 5 |
| (mean spreads) | | | | | | | | | | |
| BHL2* | 0.793 | 0.97% | 41.047 | 41.046 | 7 | 0.900 | -8.702% | 6.510 | 6.511 | 6 |
| (mean parameters) | | | | | | | | | | |
| BHL3* | 112.316 | 11122% | 23.817 | 113.742 | 13 | 46.195 | 4519% | 3.775 | 45.349 | 14 |
| (negatives to be zero) | | | | | | | | | | |
| CS1 ^ˆ | 21.221 | 2039% | 37.204 | 42.426 | 9 | 9.155 | 817% | 6.100 | 10.195 | 10 |
| (mean spreads) | | | | | | | | | | |
| CS2 ^ˆ | 2.643 | 185% | 43.125 | 43.164 | 10 | 0.999 | 1.231% | 7.139 | 7.139 | 8 |
| (mean parameters) | | | | | | | | | | |
| CS3 [‡] | 82.901 | 8180% | 19.591 | 84.112 | 12 | 34.266 | 3326% | 3.221 | 33.415 | 12 |
| (negatives to be zero) | | | | | | | | | | |
| ROLL | 93.496 | 15269% | 64.917 | 165.918 | 14 | 20.969 | 3758% | 15.947 | 40.823 | 13 |
| AR [‡] | 74.679 | 7350% | 22.047 | 76.738 | 11 | 30.438 | 2943% | 3.503 | 29.635 | 11 |
| (negatives to be zero) | | | | | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 18.052 | 1693% | 29.323 | 33.859 | 1 | 2.709 | 167% | 4.972 | 5.244 | 2 |
| Combination2 (SHL2+CS2)/2 | 18.996 | 1789% | 30.546 | 35.397 | 2 | 2.756 | 171% | 5.247 | 5.520 | 3 |
| Combination3 (CS1+BHL1)/2 | 10.988 | 1016% | 38.554 | 39.869 | 4 | 5.030 | 404% | 6.262 | 7.454 | 9 |
| Combination4 (BHL1+BHL2)/2 | 0.775 | -2.55% | 40.307 | 40.306 | 5 | 0.903 | -8.297% | 6.485 | 6.485 | 4 |

The standard deviation of daily mid-price return is 0.1897. The true spread is fixed.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^ˆThe monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The column 'Ranking' reports the rankings of the estimators according to values produced in column RMSE.

This table reports the results of the time intervals of daily and four-hours respectively. *Midstd* represents the standard deviation of mid-price returns over the relevant interval. For each time interval, there are five panels which report the summary statistics and the results of the estimators respectively. Mean indicates the average of estimated spreads over 25000 replications. Outliers of relative errors, the highest and lowest 1% of the relative estimated errors, are trimmed off before further calculation. We also report the rankings of the estimators according to RMSE.

Table 2: Simulation experiments: Comparison of the estimates over 25000 months

| | Daily (MidStd= 189.7*0.001) 20 observations per month | | | | | Four hours (MidStd =77.5*0.001) 120 observations per month | | | | |
|-------------------------------|--|--------------|---------------------------|--------|---------|---|--------------|-------------|--------|---------|
| | True spread=2 (*0.001) | | True spread/Midstd=0.0105 | | | True spread/Midstd=0.0258 | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking |
| SHL1 | 2.181 | 20.22% | 20.577 | 20.578 | 8 | 1.909 | -3.972% | 3.272 | 3.272 | 7 |
| SHL2* | 36.002 | 1680% | 10.134 | 19.620 | 3 | 5.096 | 150% | 1.951 | 2.459 | 1 |
| (negatives to be zero) | | | | | | | | | | |
| BHL1* | 1.985 | 8.91% | 20.431 | 20.431 | 6 | 1.909 | -4.022% | 3.263 | 3.263 | 5 |
| (mean spreads) | | | | | | | | | | |
| BHL2* | 2.169 | 19.27% | 20.546 | 20.547 | 7 | 1.911 | -3.887% | 3.270 | 3.270 | 6 |
| (mean parameters) | | | | | | | | | | |
| BHL3* | 113.081 | 5550% | 11.873 | 56.758 | 13 | 46.773 | 2238% | 1.912 | 22.466 | 14 |
| (negatives to be zero) | | | | | | | | | | |
| CS1 ^ˆ | 22.457 | 1032% | 18.541 | 21.217 | 9 | 10.116 | 406% | 3.069 | 5.092 | 10 |
| (mean spreads) | | | | | | | | | | |
| CS2 ^ˆ | 3.767 | 97.52% | 21.513 | 21.535 | 10 | 1.967 | -1.012% | 3.593 | 3.593 | 8 |
| (mean parameters) | | | | | | | | | | |
| CS3 [‡] | 83.778 | 4085% | 9.757 | 41.997 | 12 | 34.880 | 1644% | 1.635 | 16.518 | 12 |
| (negatives to be zero) | | | | | | | | | | |
| ROLL | 93.022 | 7581% | 32.519 | 82.487 | 14 | 21.411 | 1839% | 8.006 | 20.059 | 13 |
| AR [‡] | 74.919 | 3637% | 10.969 | 37.985 | 11 | 30.540 | 1426% | 1.756 | 14.372 | 11 |
| (negatives to be zero) | | | | | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 18.993 | 845% | 14.672 | 16.929 | 1 | 3.503 | 72.871% | 2.571 | 2.672 | 2 |
| Combination2 (SHL2+CS2)/2 | 19.884 | 889% | 15.277 | 17.674 | 2 | 3.531 | 74.460% | 2.712 | 2.812 | 3 |
| Combination3 (CS1+BHL1)/2 | 12.221 | 520% | 19.240 | 19.930 | 4 | 6.012 | 201% | 3.148 | 3.736 | 9 |
| Combination4 (BHL1+BHL2)/2 | 2.077 | 14.08% | 20.155 | 20.155 | 5 | 1.910 | -3.949% | 3.257 | 3.258 | 4 |

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^ˆThe monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (1).

Table 3: Simulation experiments: Comparison of the estimates over 25000 months

| | Daily (MidStd= 189.7*0.001) 20 observations per month | | | | | Four hours (MidStd =77.5*0.001) 120 observations per month | | | | | | | | | |
|-------------------------------|--|--------------|-------------|--------|---------|---|--------------|-------------|-------|---------|---------------------------|--------------|-------------|-------|---------|
| | True spread=6 (*0.001) | | | | | True spread/Midstd=0.0316 | | | | | True spread/Midstd=0.0775 | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking |
| SHL1 | 5.896 | 1.82% | 6.900 | 6.900 | 8 | 5.818 | -2.781% | 1.080 | 1.081 | 7 | 5.818 | -2.781% | 1.080 | 1.081 | 7 |
| SHL2* | 38.141 | 529% | 3.523 | 6.354 | 3 | 7.634 | 25.770% | 0.792 | 0.833 | 1 | 7.634 | 25.770% | 0.792 | 0.833 | 1 |
| (negatives to be zero) | | | | | | | | | | | | | | | |
| BHL1* | 5.805 | -0.41% | 6.831 | 6.831 | 6 | 5.821 | -2.709% | 1.080 | 1.080 | 5 | 5.821 | -2.709% | 1.080 | 1.080 | 5 |
| (mean spreads) | | | | | | | | | | | | | | | |
| BHL2* | 5.925 | 1.97% | 6.896 | 6.895 | 7 | 5.817 | -2.782% | 1.080 | 1.080 | 6 | 5.817 | -2.782% | 1.080 | 1.080 | 6 |
| (mean parameters) | | | | | | | | | | | | | | | |
| BHL3* | 115.153 | 1818% | 3.966 | 18.608 | 13 | 48.999 | 717% | 0.646 | 7.195 | 14 | 48.999 | 717% | 0.646 | 7.195 | 14 |
| (negatives to be zero) | | | | | | | | | | | | | | | |
| CS1 ^ˆ | 25.936 | 335% | 6.219 | 7.063 | 9 | 13.797 | 130% | 1.014 | 1.650 | 10 | 13.797 | 130% | 1.014 | 1.650 | 10 |
| (mean spreads) | | | | | | | | | | | | | | | |
| CS2 ^ˆ | 7.377 | 25.97% | 7.223 | 7.227 | 10 | 5.714 | -4.482% | 1.184 | 1.184 | 8 | 5.714 | -4.482% | 1.184 | 1.184 | 8 |
| (mean parameters) | | | | | | | | | | | | | | | |
| CS3 [‡] | 85.942 | 1331% | 3.315 | 13.713 | 12 | 37.168 | 519% | 0.561 | 5.224 | 12 | 37.168 | 519% | 0.561 | 5.224 | 12 |
| (negatives to be zero) | | | | | | | | | | | | | | | |
| ROLL | 93.559 | 2455% | 10.843 | 26.840 | 14 | 21.573 | 546% | 2.646 | 6.071 | 13 | 21.573 | 546% | 2.646 | 6.071 | 13 |
| AR [‡] | 74.826 | 1144% | 3.659 | 12.013 | 11 | 30.687 | 411% | 0.586 | 4.154 | 11 | 30.687 | 411% | 0.586 | 4.154 | 11 |
| (negatives to be zero) | | | | | | | | | | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 21.973 | 264% | 4.975 | 5.633 | 1 | 6.728 | 11.535% | 0.927 | 0.935 | 2 | 6.728 | 11.535% | 0.927 | 0.935 | 2 |
| Combination2 (SHL2+CS2)/2 | 22.759 | 277% | 5.193 | 5.888 | 2 | 6.674 | 10.668% | 0.971 | 0.977 | 3 | 6.674 | 10.668% | 0.971 | 0.977 | 3 |
| Combination3 (CS1+BHL1)/2 | 15.870 | 167% | 6.443 | 6.656 | 4 | 9.809 | 63.730% | 1.041 | 1.221 | 9 | 9.809 | 63.730% | 1.041 | 1.221 | 9 |
| Combination4 (BHL1+BHL2)/2 | 5.865 | 0.82% | 6.750 | 6.750 | 5 | 5.819 | -2.740% | 1.076 | 1.077 | 4 | 5.819 | -2.740% | 1.076 | 1.077 | 4 |

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^ˆThe monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (1).

Table 4: Simulation experiments: Comparison of the estimates over 25000 months

| | Daily (MidStd= 189.7*0.001) 20 observations per month | | | | | Four hours (MidStd =77.5*0.001) 120 observations per month | | | | |
|-------------------------------|--|--------------|-------------|--------|---------|---|--------------|-------------|-------|---------|
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking |
| SHL1 | 9.884 | 0.82% | 4.143 | 4.143 | 8 | 9.728 | -2.641% | 0.657 | 0.657 | 6 |
| SHL2* | 40.456 | 301% | 2.182 | 3.714 | 3 | 10.699 | 6.274% | 0.552 | 0.556 | 1 |
| (negatives to be zero) | | | | | | | | | | |
| BHL1* | 9.829 | -0.04% | 4.097 | 4.097 | 6 | 9.730 | -2.629% | 0.656 | 0.657 | 5 |
| (mean spreads) | | | | | | | | | | |
| BHL2* | 9.833 | 0.15% | 4.133 | 4.133 | 7 | 9.729 | -2.626% | 0.657 | 0.658 | 7 |
| (mean parameters) | | | | | | | | | | |
| BHL3* | 117.638 | 1076% | 2.385 | 11.017 | 13 | 51.223 | 412% | 0.393 | 4.141 | 14 |
| (negatives to be zero) | | | | | | | | | | |
| CS1 ^ˆ | 29.716 | 199% | 3.725 | 4.222 | 9 | 17.474 | 74.802% | 0.617 | 0.970 | 10 |
| (mean spreads) | | | | | | | | | | |
| CS2 ^ˆ | 11.212 | 14.09% | 4.313 | 4.315 | 10 | 9.454 | -5.372% | 0.718 | 0.720 | 8 |
| (mean parameters) | | | | | | | | | | |
| CS3 [‡] | 88.401 | 783% | 2.014 | 8.085 | 12 | 39.539 | 295% | 0.349 | 2.974 | 12 |
| (negatives to be zero) | | | | | | | | | | |
| ROLL | 93.914 | 1433% | 6.524 | 15.746 | 14 | 21.875 | 293% | 1.611 | 3.343 | 13 |
| AR [‡] | 75.142 | 650% | 2.200 | 6.860 | 11 | 30.997 | 210% | 0.355 | 2.128 | 11 |
| (negatives to be zero) | | | | | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 25.142 | 150% | 3.018 | 3.371 | 1 | 10.214 | 1.822% | 0.601 | 0.601 | 2 |
| Combination2 (SHL2+CS2)/2 | 25.834 | 157% | 3.140 | 3.512 | 2 | 10.076 | 0.455% | 0.626 | 0.626 | 3 |
| Combination3 (CS1+BHL1)/2 | 19.772 | 99.33% | 3.862 | 3.987 | 4 | 13.602 | 36.083% | 0.633 | 0.729 | 9 |
| Combination4 (BHL1+BHL2)/2 | 9.831 | 0.06% | 4.048 | 4.048 | 5 | 9.729 | -2.628% | 0.655 | 0.655 | 4 |

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^ˆThe monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (1).

Table 5: Simulation experiments: Comparison of the estimates over 25000 months

| | Daily (MidStd= 189.7*0.001) 20 observations per month | | | | | Four hours (MidStd =77.5*0.001) 120 observations per month | | | | |
|-------------------------------|--|--------------|--------------------------|-------|---------|---|--------------|-------------|-------|---------|
| | True spread=30 (*0.001) | | True spread/Midstd=0.158 | | | True spread/Midstd=0.387 | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking |
| SHL1 | 29.607 | -0.61% | 1.387 | 1.387 | 8 | 29.809 | -0.603% | 0.218 | 0.218 | 5 |
| SHL2* | 52.870 | 74.94% | 0.833 | 1.121 | 1 | 29.818 | -0.588% | 0.218 | 0.218 | 4 |
| (negatives to be zero) | | | | | | | | | | |
| BHL1* | 29.567 | -0.92% | 1.375 | 1.375 | 6 | 29.820 | -0.573% | 0.218 | 0.218 | 3 |
| (mean spreads) | | | | | | | | | | |
| BHL2* | 29.595 | -0.74% | 1.385 | 1.385 | 7 | 29.809 | -0.607% | 0.218 | 0.218 | 6 |
| (mean parameters) | | | | | | | | | | |
| BHL3* | 128.947 | 330% | 0.826 | 3.398 | 13 | 63.517 | 112% | 0.142 | 1.126 | 14 |
| (negatives to be zero) | | | | | | | | | | |
| CS1^ | 48.244 | 61.34% | 1.249 | 1.391 | 9 | 36.619 | 22.091% | 0.206 | 0.302 | 11 |
| (mean spreads) | | | | | | | | | | |
| CS2^ | 30.054 | 0.83% | 1.439 | 1.439 | 10 | 29.055 | -3.116% | 0.237 | 0.239 | 10 |
| (mean parameters) | | | | | | | | | | |
| CS3‡ | 100.202 | 234% | 0.724 | 2.447 | 12 | 52.595 | 75.295% | 0.135 | 0.765 | 13 |
| (negatives to be zero) | | | | | | | | | | |
| ROLL | 95.550 | 416% | 2.206 | 4.706 | 14 | 30.286 | 47.501% | 0.582 | 0.751 | 12 |
| AR‡ | 76.747 | 155% | 0.750 | 1.724 | 11 | 35.313 | 17.669% | 0.131 | 0.220 | 7 |
| (negatives to be zero) | | | | | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 41.218 | 36.99% | 1.067 | 1.129 | 2 | 29.819 | -0.581% | 0.217 | 0.217 | 1 |
| Combination2 (SHL2+CS2)/2 | 41.462 | 37.88% | 1.103 | 1.166 | 3 | 29.437 | -1.856% | 0.226 | 0.226 | 8 |
| Combination3 (CS1+BHL1)/2 | 38.905 | 30.19% | 1.296 | 1.331 | 4 | 33.219 | 10.758% | 0.211 | 0.237 | 9 |
| Combination4 (BHL1+BHL2)/2 | 29.581 | -0.83% | 1.358 | 1.358 | 5 | 29.814 | -0.590% | 0.218 | 0.218 | 2 |

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (1).

Table 6: Simulation experiments: Cross-sectional properties of the estimates

| | Mean true spread=11.02 (*0.001) Range from 2.00 to 20.00 (*0.001) | Daily (MidStd= 189.7*0.001) Truespread/Midstd=0.0581 75000 months 20 observations per month | | | | |
|-------------------------------|--|--|-------------|--------|---------|--------------|
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
| SHL1 | 10.461 | -2.53% | 5.612 | 5.612 | 7 | 0.113 |
| SHL2* | 40.867 | 378% | 4.006 | 5.506 | 5 | 0.127 |
| (negatives to be zero) | | | | | | |
| BHL1* | 10.429 | -3.79% | 5.560 | 5.560 | 6 | 0.115 |
| (mean spreads) | | | | | | |
| BHL2* | 10.438 | -2.87% | 5.618 | 5.618 | 8 | 0.114 |
| (mean parameters) | | | | | | |
| BHL3* | 117.774 | 1343% | 10.005 | 16.748 | 13 | 0.112 |
| (negatives to be zero) | | | | | | |
| CS1 [^] | 30.265 | 242% | 5.339 | 5.863 | 10 | 0.119 |
| (mean spreads) | | | | | | |
| CS2 [^] | 11.820 | 14.74% | 5.849 | 5.851 | 9 | 0.107 |
| (mean parameters) | | | | | | |
| CS3 [‡] | 88.586 | 979.10% | 7.484 | 12.324 | 12 | 0.136 |
| (negatives to be zero) | | | | | | |
| ROLL | 93.103 | 1759% | 14.736 | 22.950 | 14 | 0.003 |
| AR [‡] | 74.796 | 825.24% | 6.948 | 10.788 | 11 | 0.008 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 25.648 | 186% | 4.286 | 4.673 | 1 | 0.124 |
| Combination2 (SHL2+CS2)/2 | 26.343 | 195% | 4.462 | 4.871 | 2 | 0.117 |
| Combination3 (CS1+BHL1)/2 | 20.347 | 119% | 5.314 | 5.445 | 3 | 0.118 |
| Combination4 (BHL1+BHL2)/2 | 10.434 | -3.38% | 5.499 | 5.499 | 4 | 0.116 |

The standard deviation of daily mid-price return is 0.1897

Mean indicates the average of estimated spreads over 75000 months. The true spread changes every month ranging from 0.002 to 0.02.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

[^]The monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The column 'Ranking' reports the rankings of the estimators according to values produced in column RMSE.

This table reports the results of the time interval of daily. *Midstd* represents the standard deviation of mid-price returns over the relevant interval. For each time interval, there are five panels which report the summary statistics and the results of the estimators respectively. Outliers of relative errors, the highest and lowest 1% of the relative estimated errors, are trimmed off before further calculation.

Table 7: Simulation experiments: Cross-sectional properties of the estimates

| | Mean true spread=3.57 (*0.001) | Daily (MidStd= 189.7*0.001) Truespread/Midstd=0.0188 | | | | |
|-------------------------------|----------------------------------|--|-------------|--------|---------|--------------|
| | Range from 2.00 to 5.13 (*0.001) | 15000 months 20 observations per month | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
| SHL1 | 3.403 | 3.51% | 12.511 | 12.511 | 7 | 0.010 |
| SHL2* | 36.725 | 983% | 6.814 | 11.957 | 3 | 0.010 |
| (negatives to be zero) | | | | | | |
| BHL1* | 3.253 | -3.09% | 12.375 | 12.375 | 6 | 0.013 |
| (mean spreads) | | | | | | |
| BHL2* | 3.494 | 4.31% | 12.543 | 12.543 | 8 | 0.015 |
| (mean parameters) | | | | | | |
| BHL3* | 113.768 | 3299.10% | 11.395 | 34.903 | 13 | 0.019 |
| (negatives to be zero) | | | | | | |
| CS1^ | 23.663 | 605.75% | 11.388 | 12.899 | 9 | 0.017 |
| (mean spreads) | | | | | | |
| CS2^ | 4.996 | 47.54% | 13.054 | 13.062 | 10 | 0.018 |
| (mean parameters) | | | | | | |
| CS3‡ | 84.477 | 2423% | 8.860 | 25.801 | 12 | 0.012 |
| (negatives to be zero) | | | | | | |
| ROLL | 93.085 | 4448% | 22.687 | 49.929 | 14 | 0.003 |
| AR‡ | 74.800 | 2133% | 8.872 | 23.101 | 11 | -0.005 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 19.989 | 489% | 9.053 | 10.291 | 1 | 0.013 |
| Combination2 (SHL2+CS2)/2 | 20.861 | 515% | 9.429 | 10.742 | 2 | 0.016 |
| Combination3 (CS1+BHL1)/2 | 13.458 | 301% | 11.701 | 12.082 | 4 | 0.015 |
| Combination4 (BHL1+BHL2)/2 | 3.373 | 0.36% | 12.252 | 12.251 | 5 | 0.015 |

Mean indicates the average of estimated spreads over 15000 months. The true spread changes every month ranging from 0.002 to 0.00513.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (6).

Table 8: Simulation experiments: Cross-sectional properties of the estimates

| | Mean true spread=6.71 (*0.001) | Daily (MidStd= 189.7*0.001) Truespread/Midstd=0.0354 | | | | |
|-------------------------------|----------------------------------|--|-------------|--------|---------|--------------|
| | Range from 5.13 to 8.29 (*0.001) | 15000 months 20 observations per month | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
| SHL1 | 6.533 | 0.62% | 6.311 | 6.310 | 8 | 0.022 |
| SHL2* | 38.490 | 476% | 3.296 | 5.786 | 3 | 0.020 |
| (negatives to be zero) | | | | | | |
| BHL1* | 6.521 | -1.06% | 6.268 | 6.268 | 6 | 0.026 |
| (mean spreads) | | | | | | |
| BHL2* | 6.458 | -0.96% | 6.294 | 6.294 | 7 | 0.024 |
| (mean parameters) | | | | | | |
| BHL3* | 115.408 | 1646% | 4.232 | 16.993 | 13 | 0.024 |
| (negatives to zero) | | | | | | |
| CS1^ | 26.490 | 302% | 5.709 | 6.457 | 9 | 0.025 |
| (mean spreads) | | | | | | |
| CS2^ | 7.966 | 21.36% | 6.610 | 6.613 | 10 | 0.023 |
| (mean parameters) | | | | | | |
| CS3‡ | 86.192 | 1203% | 3.458 | 12.518 | 12 | 0.022 |
| (negatives to zero) | | | | | | |
| ROLL | 93.129 | 2224% | 10.244 | 24.486 | 14 | 0.001 |
| AR‡ | 74.634 | 1028% | 3.641 | 10.905 | 11 | 0.005 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 22.505 | 237% | 4.573 | 5.152 | 1 | 0.025 |
| Combination2 (SHL2+CS2)/2 | 23.228 | 248% | 4.765 | 5.374 | 2 | 0.022 |
| Combination3 (CS1+BHL1)/2 | 16.505 | 150% | 5.908 | 6.095 | 4 | 0.026 |
| Combination4 (BHL1+BHL2)/2 | 6.489 | -0.96% | 6.177 | 6.177 | 5 | 0.026 |

Mean indicates the average of estimated spreads over 15000 months. The true spread changes every month ranging from 0.00513 to 0.00829.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (6).

Table 9: Simulation experiments: Cross-sectional properties of the estimates

| Mean true spread=9.86 (*0.001) | | Daily (MidStd= 189.7*0.001) Truespread/Midstd=0.0520 | | | | |
|-----------------------------------|------------|--|-------------|--------|---------|--------------|
| Range from 8.29 to 11.44 (*0.001) | | 15000 months 20 observations per month | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
| SHL1 | 8.527 | -11.13% | 4.249 | 4.250 | 7 | 0.019 |
| SHL2* | 39.773 | 302% | 2.218 | 3.744 | 3 | 0.021 |
| (negatives to be zero) | | | | | | |
| BHL1* | 8.584 | -10.81% | 4.222 | 4.223 | 6 | 0.018 |
| (mean spreads) | | | | | | |
| BHL2* | 8.323 | -13.21% | 4.251 | 4.253 | 8 | 0.018 |
| (mean parameters) | | | | | | |
| BHL3* | 116.896 | 1093% | 2.644 | 11.244 | 13 | 0.020 |
| (negatives to be zero) | | | | | | |
| CS1 [^] | 28.435 | 192% | 3.836 | 4.289 | 9 | 0.018 |
| (mean spreads) | | | | | | |
| CS2 [^] | 9.763 | 1.38% | 4.445 | 4.445 | 10 | 0.016 |
| (mean parameters) | | | | | | |
| CS3 [‡] | 87.575 | 793% | 2.191 | 8.230 | 12 | 0.021 |
| (negatives to be zero) | | | | | | |
| ROLL | 92.115 | 1447% | 6.774 | 15.979 | 14 | -0.002 |
| AR [‡] | 74.396 | 658% | 2.345 | 6.987 | 11 | 0.004 |
| (negatives to be zero) | | | | | | |
| Combination1 | 24.178 | 145% | 3.090 | 3.415 | 1 | 0.020 |
| (BHL1+SHL2)/2 | | | | | | |
| Combination2 | 24.768 | 152% | 3.214 | 3.554 | 2 | 0.018 |
| (SHL2+CS2)/2 | | | | | | |
| Combination3 | 18.509 | 90.50% | 3.978 | 4.079 | 4 | 0.018 |
| (CS1+BHL1)/2 | | | | | | |
| Combination4 | 8.453 | -12.01% | 4.169 | 4.171 | 5 | 0.018 |
| (BHL1+BHL2)/2 | | | | | | |

Mean indicates the average of estimated spreads over 15000 months. The true spread changes every month ranging from 0.00829 to 0.0114.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

[^]The monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (6).

Table 10: Simulation experiments: Cross-sectional properties of the estimates

| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|-------------------------------|------------|--------------|-------------|--------|---------|--------------|
| SHL1 | 12.128 | -4.66% | 3.218 | 3.218 | 8 | 0.020 |
| SHL2* | 41.823 | 220% | 1.730 | 2.798 | 3 | 0.025 |
| (negatives to be zero) | | | | | | |
| BHL1* | 12.094 | -5.31% | 3.198 | 3.198 | 6 | 0.018 |
| (mean spreads) | | | | | | |
| BHL2* | 12.081 | -5.39% | 3.217 | 3.217 | 7 | 0.020 |
| (mean parameters) | | | | | | |
| BHL3* | 118.783 | 818% | 1.939 | 8.403 | 13 | 0.028 |
| (negatives to be zero) | | | | | | |
| CS1 [^] | 31.829 | 147% | 2.907 | 3.258 | 9 | 0.020 |
| (mean spreads) | | | | | | |
| CS2 [^] | 13.532 | 5.84% | 3.363 | 3.364 | 10 | 0.018 |
| (mean parameters) | | | | | | |
| CS3 [‡] | 89.523 | 591% | 1.642 | 6.137 | 12 | 0.027 |
| (negatives to be zero) | | | | | | |
| ROLL | 92.732 | 1086% | 5.084 | 11.992 | 14 | 0.010 |
| AR [‡] | 74.802 | 477% | 1.745 | 5.080 | 11 | 0.012 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 26.958 | 107% | 2.370 | 2.601 | 1 | 0.022 |
| Combination2 (SHL2+CS2)/2 | 27.678 | 113% | 2.463 | 2.709 | 2 | 0.021 |
| Combination3 (CS1+BHL1)/2 | 21.961 | 70.83% | 3.013 | 3.095 | 4 | 0.020 |
| Combination4 (BHL1+BHL2)/2 | 12.087 | -5.39% | 3.156 | 3.156 | 5 | 0.020 |

Mean indicates the average of estimated spreads over 15000 months. The true spread changes every month ranging from 0.0114 to 0.0146.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

[^]The monthly CS estimates can be calculated using the two methods described in note *.

[‡] Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (6).

Table 11: Simulation experiments: Cross-sectional properties of the estimates

| | Mean true spread=16.10 (*0.001) | Daily (MidStd= 189.7*0.001) Truespread/Midstd=0.0849 | | | | |
|-------------------------------|------------------------------------|--|-------------|-------|---------|--------------|
| | Range from 14.55 to 17.65 (*0.001) | 15000 months 20 observations per month | | | | |
| | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
| SHL1 | 15.941 | 0.68% | 2.590 | 2.590 | 7 | 0.010 |
| SHL2* | 44.044 | 172% | 1.425 | 2.232 | 3 | 0.008 |
| (negatives to be zero) | | | | | | |
| BHL1* | 15.980 | 0.71% | 2.577 | 2.577 | 6 | 0.013 |
| (mean spreads) | | | | | | |
| BHL2* | 15.981 | 0.82% | 2.596 | 2.596 | 8 | 0.011 |
| (mean parameters) | | | | | | |
| BHL3* | 120.740 | 652% | 1.559 | 6.699 | 13 | 0.012 |
| (negatives to be zero) | | | | | | |
| CS1^ | 35.399 | 122% | 2.346 | 2.642 | 9 | 0.009 |
| (mean spreads) | | | | | | |
| CS2^ | 17.234 | 8.60% | 2.718 | 2.720 | 10 | 0.009 |
| (mean parameters) | | | | | | |
| CS3‡ | 91.672 | 470% | 1.316 | 4.885 | 12 | 0.011 |
| (negatives to be zero) | | | | | | |
| ROLL | 93.978 | 862% | 4.100 | 9.543 | 14 | -0.010 |
| AR‡ | 74.862 | 365% | 1.409 | 3.915 | 11 | -0.010 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 30.012 | 86.30% | 1.928 | 2.112 | 1 | 0.012 |
| Combination2 (SHL2+CS2)/2 | 30.639 | 90.25% | 2.008 | 2.201 | 2 | 0.009 |
| Combination3 (CS1+BHL1)/2 | 25.689 | 61.12% | 2.431 | 2.507 | 4 | 0.011 |
| Combination4 (BHL1+BHL2)/2 | 15.981 | 0.78% | 2.546 | 2.546 | 5 | 0.012 |

Mean indicates the average of estimated spreads over 15000 months. The true spread changes every month ranging from 0.0146 to 0.0177.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

The other settings are the same as Table (6).

Table 12: Simulation experiments: Average Ranking according to RMSE

| | All cases according to Tables 1 to 11 | Fixed spread cases according to Tables 1 to 5 | Cross-sectional cases according to Tables 6 to 11 |
|-------------------------------|--|--|--|
| SHL1 | 7.3 | 7.2 | 7.4 |
| SHL2* (negatives to be zero) | 2.4 | 2.1 | 3 |
| BHL1 (mean spreads) * | 5.5 | 5.3 | 6 |
| BHL2 (mean parameters) * | 6.9 | 6.6 | 7.6 |
| BHL3 (negatives to be zero) * | 13.3 | 13.5 | 13 |
| CS1^(mean spreads) | 9.4 | 9.6 | 9 |
| CS2^(mean parameters) | 9.5 | 9.2 | 10 |
| CS3‡ (negatives to be zero) | 12.1 | 12.1 | 12 |
| ROLL | 13.6 | 13.4 | 14 |
| AR‡ (negatives to be zero) | 10.7 | 10.6 | 11 |
| Combination1 (BHL1+SHL2)/2 | 1.3 | 1.5 | 1 |
| Combination2 (SHL2+CS2)/2 | 2.7 | 3.1 | 2 |
| Combination3 (CS1+BHL1)/2 | 5.7 | 6.5 | 4 |
| Combination4 (BHL1+BHL2)/2 | 4.5 | 4.3 | 5 |

This table reports the average ranking of the estimators and the combinations in simulations experiments according to the columns of ranking in Tables 1 to 11.

Table 13: DataStream 22 Currency pairs from 1990 to 2016 pooled

| Mean closing spread = 14.24 (*0.001) | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|---|------------|--------------|-------------|-------|---------|--------------|
| SHL1 | 7.392 | -44.49% | 2.356 | 2.397 | 9 | 0.043 |
| SHL2* | 29.880 | 90.49% | 1.398 | 1.665 | 2 | 0.726 |
| (negatives to be zero) | | | | | | |
| BHL1* | 9.044 | -37.01% | 2.243 | 2.274 | 6 | 0.076 |
| (mean spreads) | | | | | | |
| BHL2* | 8.636 | -38.69% | 2.368 | 2.399 | 10 | 0.069 |
| (mean parameters) | | | | | | |
| BHL3* | 94.082 | 505% | 3.835 | 6.343 | 13 | 0.869 |
| (negatives to be zero) | | | | | | |
| CS1^ | 29.406 | 91.16% | 1.993 | 2.192 | 5 | 0.628 |
| (mean spreads) | | | | | | |
| CS2^ | 16.944 | 17.05% | 2.351 | 2.357 | 8 | 0.251 |
| (mean parameters) | | | | | | |
| CS3‡ | 60.026 | 313% | 2.488 | 4.000 | 12 | 0.876 |
| (negatives to be zero) | | | | | | |
| ROLL | 64.439 | 580% | 5.548 | 8.031 | 14 | 0.591 |
| AR‡ | 54.120 | 253% | 2.441 | 3.517 | 11 | 0.832 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 19.462 | 26.73% | 1.637 | 1.659 | 1 | 0.423 |
| Combination2 (SHL2+CS2)/2 | 23.412 | 53.77% | 1.744 | 1.825 | 3 | 0.505 |
| Combination3 (CS1+BHL1)/2 | 19.225 | 26.71% | 2.033 | 2.050 | 4 | 0.367 |
| Combination4 (BHL1+BHL2)/2 | 8.838 | -37.79% | 2.249 | 2.280 | 7 | 0.074 |

The results above refer to the testing of the following currency pairs: AUD/USD, CAD/USD, CHF/USD, DKK/USD, EUR/USD, GBP/USD, NOK/USD, SEK/USD, USD/JPY, USD/NZD, AUD/EUR, CAD/GBP, CHF/EUR, DKK/GBP, EUR/GBP, EUR/JPY, GBP/AUD, GBP/JPY, NOK/EUR, NOK/GBP, SEK/EUR, SEK/GBP. Daily closing spreads are used to calculate the monthly benchmark bid-ask spread.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

Table 14: Currency-by-currency Average Ranking DataStream from 1990 to 2016

| | All cases |
|-------------------------|-----------|
| SHL1 | 8 |
| SHL2 | 1.7 |
| (negatives to be zero)* | |
| BHL1 | 6 |
| (mean spreads)* | |
| BHL2 | 8.3 |
| (mean parameters)* | |
| BHL3 | 13.1 |
| (negatives to be zero)* | |
| CS1 | 7.3 |
| (mean spreads)^ | |
| CS2 | 8.2 |
| (mean parameters)^ | |
| CS3‡ | 12 |
| (negatives to be zero) | |
| ROLL | 13.9 |
| AR‡ | 10 |
| (negatives to be zero) | |
| Combination1 | 1.5 |
| (SHL2+BHL1)/2 | |
| Combination2 | 3.8 |
| (SHL2+CS2)/2 | |
| Combination3 | 5.5 |
| (CS1+BHL1)/2 | |
| Combination4 | 5.6 |
| (BHL1+BHL2)/2 | |

This table reports the average ranking of the estimators and the combinations for 22 currency pairs according to RMSE.

Currency pairs used in this table are listed in Table 13

Table 15: Currency-by-currency average correlation DataStream from 1990 to 2016

| | Time series correlation | Cross-sectional correlation |
|------------------------------|-------------------------|-----------------------------|
| SHL1 | 0.091 | 0.147 |
| SHL2* (negatives to be zero) | 0.198 | 0.864 |
| BHL1* (mean spreads) | 0.074 | 0.176 |
| BHL2* (mean parameters) | 0.092 | 0.179 |
| BHL3* (negatives to be zero) | 0.271 | 0.951 |
| CS1^(mean spreads) | 0.158 | 0.636 |
| CS2^(mean parameters) | 0.125 | 0.297 |
| CS3‡ (negatives to be zero) | 0.265 | 0.948 |
| ROLL | 0.108 | 0.543 |
| AR‡(negatives to be zero) | 0.254 | 0.948 |
| Combination1(BHL1+SHL2)/2 | 0.127 | 0.443 |
| Combination2(SHL2+CS2)/2 | 0.158 | 0.518 |
| Combination3(CS1+BHL1)/2 | 0.118 | 0.4 |
| Combination4(BHL1+BHL2)/2 | 0.085 | 0.187 |

This table reports the average time-series and cross-sectional correlation of the estimators and the combinations.

Currency pairs used in this table are listed in Table 13

Daily closing spreads are used to calculate the monthly benchmark bid-ask spread.

Highest two correlation coefficients are made bold.

Table 16: Hotspot 12 Currency pairs from 2015.12 to 2016.8

| Effective Spread = 3.012 (*0.001) | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|--------------------------------------|------------|--------------|-------------|--------|---------|--------------|
| SHL1 | -41.563 | -1632% | 24.020 | 28.944 | 10 | -0.477 |
| SHL2* | 34.085 | 754% | 7.659 | 10.718 | 1 | 0.241 |
| (negatives to be zero) | | | | | | |
| BHL1* | -31.939 | -1204% | 21.287 | 24.364 | 7 | -0.500 |
| (mean spreads) | | | | | | |
| BHL2* (mean parameters) | -38.304 | -1618% | 25.328 | 29.955 | 11 | -0.436 |
| BHL3* | 169.091 | 5768% | 33.146 | 66.443 | 14 | 0.716 |
| (negatives to be zero) | | | | | | |
| CS1^ | 26.925 | 876% | 10.814 | 13.874 | 5 | 0.088 |
| (mean spreads) | | | | | | |
| CS2^ | 12.673 | 492% | 14.244 | 15.004 | 6 | -0.086 |
| (mean parameters) | | | | | | |
| CS3‡ | 74.603 | 2515% | 12.350 | 27.990 | 9 | 0.567 |
| (negatives to be zero) | | | | | | |
| ROLL | 99.593 | 4559% | 28.853 | 53.834 | 12 | 0.596 |
| AR‡ | 159.270 | 5191% | 25.001 | 57.561 | 13 | 0.732 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | 1.073 | -229% | 11.608 | 11.777 | 3 | -0.218 |
| Combination2 (SHL2+CS2)/2 | 23.379 | 627% | 9.579 | 11.412 | 2 | 0.078 |
| Combination3 (CS1+BHL1)/2 | -2.507 | -163% | 13.433 | 13.467 | 4 | -0.296 |
| Combination4 (BHL1+BHL2)/2 | -35.121 | -1411% | 23.044 | 26.924 | 8 | -0.476 |

The results above refer to the testing of the following currency pairs: AUD/USD, EUR/GBP, EUR/JPY, EUR/SEK, EUR/USD, GBP/USD, NZD/USD, USD/CAD, USD/CHF, USD/CHF, USD/JPY, USD/MXN, USD/ZAR. Tick by tick transaction and quoted data are used to generate the monthly effective spread.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

Table 17: Currency-by-currency average correlation Hotspot from 2015.12 to 2016.8

| | Time series correlation | Cross-sectional correlation |
|-------------------------------|-------------------------|-----------------------------|
| SHL1 | -0.409 | -0.325 |
| SHL2 (negatives to be zero) * | 0.035 | 0.636 |
| BHL1 (mean spreads) * | -0.292 | -0.501 |
| BHL2 (mean parameters) * | -0.298 | -0.331 |
| BHL3 (negatives to be zero) * | 0.514 | 0.772 |
| CS1 (mean spreads) ^ | 0.058 | 0.338 |
| CS2 (mean parameters) ^ | 0.042 | -0.133 |
| CS3‡ (negatives to be zero) | 0.321 | 0.739 |
| ROLL | 0.239 | 0.575 |
| AR‡ (negatives to be zero) | 0.480 | 0.800 |
| Combination1 (BHL1+SHL2)/2 | -0.235 | -0.156 |
| Combination2 (SHL2+CS2)/2 | 0.081 | 0.192 |
| Combination3 (CS1+BHL1)/2 | -0.194 | -0.305 |
| Combination4 (BHL1+BHL2)/2 | -0.297 | -0.383 |

This table reports the average time-series and cross-sectional correlations of the estimators and the combinations.

Currency pairs used in this table are listed in Table 16

Tick by tick transaction and quoted data are used to generate the monthly effective spread.

Highest two correlation coefficients are made bold.

Table 18: TAQ (Quoted Spread) S&P 1500 2014.01-2014.12

| Quoted Spread = 310.2 (*0.001) | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|-----------------------------------|------------|--------------|-------------|-------|---------|--------------|
| SHL1 | -264 | -356% | 3.758 | 5.177 | 11 | -0.086 |
| SHL2* | 70 | -59% | 0.512 | 0.779 | 1 | 0.743 |
| (negatives to be zero) | | | | | | |
| BHL1* | -277 | -367% | 3.764 | 5.259 | 12 | -0.083 |
| (mean spreads) | | | | | | |
| BHL2* | -256 | -351% | 3.751 | 5.135 | 9 | -0.082 |
| (mean parameters) | | | | | | |
| BHL3* | 362 | 193% | 2.648 | 3.274 | 7 | 0.779 |
| (negatives to be zero) | | | | | | |
| CS1^ | -44 | -211% | 2.277 | 3.106 | 5 | 0.301 |
| (mean spreads) | | | | | | |
| CS2^ | -236 | -375% | 4.332 | 5.732 | 13 | 0.009 |
| (mean parameters) | | | | | | |
| CS3‡ | 222 | 104% | 1.849 | 2.119 | 2 | 0.527 |
| (negatives to be zero) | | | | | | |
| ROLL | 426 | 466% | 6.109 | 7.683 | 14 | 0.276 |
| AR‡ | 306 | 174% | 2.748 | 3.255 | 6 | 0.679 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | -104 | -212% | 1.903 | 2.852 | 3 | 0.030 |
| Combination2 (SHL2+CS2)/2 | -83 | -217% | 2.208 | 3.094 | 4 | 0.303 |
| Combination3 (CS1+BHL1)/2 | -161 | -290% | 2.985 | 4.164 | 8 | -0.039 |
| Combination4 (BHL1+BHL2)/2 | -267 | -359% | 3.728 | 5.175 | 10 | -0.083 |

Tick by tick quoted data are used to generate the monthly quoted spread.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

Table 19: TAQ (Quoted Spread) S&P 500 2014.01-2014.12

| Quoted Spread = 202.3 (*0.001) | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|-----------------------------------|------------|--------------|-------------|--------|---------|--------------|
| SHL1 | -462 | -632% | 5.317 | 8.256 | 11 | -0.273 |
| SHL2* | 55 | -40% | 0.743 | 0.843 | 1 | 0.527 |
| (negatives to be zero) | | | | | | |
| BHL1* | -468 | -646% | 5.264 | 8.329 | 12 | -0.292 |
| (mean spreads) | | | | | | |
| BHL2* | -451 | -624% | 5.358 | 8.227 | 9 | -0.286 |
| (mean parameters) | | | | | | |
| BHL3* | 416 | 415% | 3.253 | 5.270 | 6 | 0.858 |
| (negatives to be zero) | | | | | | |
| CS1^ | -106 | -348% | 3.322 | 4.808 | 5 | 0.204 |
| (mean spreads) | | | | | | |
| CS2^ | -370 | -667% | 6.230 | 9.127 | 13 | -0.275 |
| (mean parameters) | | | | | | |
| CS3‡ | 241 | 256% | 2.287 | 3.431 | 2 | 0.623 |
| (negatives to be zero) | | | | | | |
| ROLL | 569 | 981% | 8.223 | 12.797 | 14 | 0.368 |
| AR‡ | 402 | 413% | 3.438 | 5.375 | 7 | 0.814 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | -207 | -342% | 2.727 | 4.373 | 3 | -0.256 |
| Combination2 (SHL2+CS2)/2 | -157 | -353% | 3.236 | 4.788 | 4 | -0.097 |
| Combination3 (CS1+BHL1)/2 | -287 | -499% | 4.218 | 6.532 | 8 | -0.269 |
| Combination4 (BHL1+BHL2)/2 | -460 | -635% | 5.261 | 8.244 | 10 | -0.290 |

Tick by tick quoted data are used to generate the monthly quoted spread.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

Table 20: TAQ (Quoted Spread) S&P 400 2014.01-2014.12

| Quoted Spread = 365 (*0.001) | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|---------------------------------|------------|--------------|-------------|-------|---------|--------------|
| SHL1 | -225 | -304% | 2.892 | 4.199 | 10 | -0.011 |
| SHL2* | 78 | -61% | 0.454 | 0.757 | 1 | 0.829 |
| (negatives to be zero) | | | | | | |
| BHL1* | -240 | -317% | 2.911 | 4.302 | 12 | 0.025 |
| (mean spreads) | | | | | | |
| BHL2* | -213 | -298% | 2.905 | 4.165 | 9 | 0.019 |
| (mean parameters) | | | | | | |
| BHL3* | 373 | 155% | 1.987 | 2.519 | 5 | 0.817 |
| (negatives to be zero) | | | | | | |
| CS1^ | -45 | -189% | 1.828 | 2.626 | 7 | 0.369 |
| (mean spreads) | | | | | | |
| CS2^ | -225 | -321% | 3.445 | 4.710 | 13 | 0.101 |
| (mean parameters) | | | | | | |
| CS3‡ | 219 | 79% | 1.398 | 1.603 | 2 | 0.584 |
| (negatives to be zero) | | | | | | |
| ROLL | 393 | 372% | 4.561 | 5.885 | 14 | 0.459 |
| AR‡ | 295 | 129% | 1.984 | 2.367 | 3 | 0.867 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | -81 | -188% | 1.497 | 2.406 | 4 | 0.292 |
| Combination2 (SHL2+CS2)/2 | -73 | -191% | 1.782 | 2.611 | 6 | 0.525 |
| Combination3 (CS1+BHL1)/2 | -142 | -253% | 2.328 | 3.441 | 8 | 0.093 |
| Combination4 (BHL1+BHL2)/2 | -226 | -308% | 2.884 | 4.215 | 11 | 0.022 |

Tick by tick quoted data are used to generate the monthly quoted spread.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

Table 21: TAQ (Quoted Spread) S&P 600 2014.01-2014.12

| Quoted Spread = 364 (*0.001) | Mean*0.001 | Rel-Err-Mean | Rel-Err-Std | RMSE | Ranking | Correlation |
|---------------------------------|------------|--------------|-------------|-------|---------|--------------|
| SHL1 | -123 | -176% | 1.324 | 2.203 | 10 | -0.024 |
| SHL2* | 76 | -72% | 0.290 | 0.776 | 2 | 0.614 |
| (negatives to be zero) | | | | | | |
| BHL1* | -139 | -184% | 1.351 | 2.282 | 12 | -0.040 |
| (mean spreads) | | | | | | |
| BHL2* | -120 | -174% | 1.315 | 2.181 | 9 | -0.042 |
| (mean parameters) | | | | | | |
| BHL3* (negatives to be zero) | 307 | 40% | 1.033 | 1.107 | 4 | 0.745 |
| CS1^ | 9 | -120% | 0.837 | 1.462 | 6 | 0.352 |
| (mean spreads) | | | | | | |
| CS2^ | -129 | -184% | 1.644 | 2.468 | 13 | 0.056 |
| (mean parameters) | | | | | | |
| CS3‡ | 207 | -1% | 0.733 | 0.733 | 1 | 0.576 |
| (negatives to be zero) | | | | | | |
| ROLL | 324 | 140% | 2.236 | 2.639 | 14 | 0.308 |
| AR‡ | 232 | 14% | 0.975 | 0.985 | 3 | 0.691 |
| (negatives to be zero) | | | | | | |
| Combination1 (BHL1+SHL2)/2 | -31 | -128% | 0.712 | 1.462 | 5 | 0.150 |
| Combination2 (SHL2+CS2)/2 | -27 | -128% | 0.869 | 1.547 | 7 | 0.231 |
| Combination3 (CS1+BHL1)/2 | -65 | -152% | 1.068 | 1.859 | 8 | 0.093 |
| Combination4 (BHL1+BHL2)/2 | -129 | -179% | 1.320 | 2.223 | 11 | -0.041 |

Tick by tick quoted data are used to generate the monthly quoted spread.

*In the instance where the SHL estimate in a trail is a negative value, we set all negative estimated spreads in a trail to zero.

*The BHL estimates can be calculated using two methods: (1) calculate the two-day interval spread for one equity finding the monthly mean for the spread (reported as 'BHL1 mean spreads' in the table above); 2) calculate the average daily and two day interval range each month and then calculate the spread (reported as 'BHL2 mean parameters' above).

^The monthly CS estimates can be calculated using the two methods described in note *.

‡ Estimates are calculated in a manner similar to that described in notes *.

Table 22: Average Ranking S&P 2014.01-2014.12

| | All cases |
|------------------------------|-----------|
| SHL1 | 10.3 |
| SHL2 (negatives to be zero)* | 1.3 |
| BHL1 (mean spreads)* | 12 |
| BHL2 (mean parameters)* | 9 |
| BHL3 (negatives to be zero)* | 5 |
| CS1 (mean spreads)^ | 6 |
| CS2 (mean parameters)^ | 13 |
| CS3‡ (negatives to be zero) | 1.7 |
| ROLL | 14 |
| AR‡ (negatives to be zero) | 4.3 |
| Combination1 (SHL2+BHL1)/2 | 4.0 |
| Combination2 (SHL2+CS2)/2 | 5.7 |
| Combination3 (CS1+BHL1)/2 | 8.0 |
| Combination4 (BHL1+BHL2)/2 | 10.7 |

This table reports the average ranking of the estimators and the combinations according to the columns of ranking in Tables 19 to 21.

Table 23: Stock-by-stock average correlation S&P 1500 2014.01-2014.12

| | Time series correlation | Cross-sectional correlation |
|-------------------------------|-------------------------|-----------------------------|
| SHL1 | -0.052 | -0.105 |
| SHL2 (negatives to be zero) * | 0.088 | 0.749 |
| BHL1 (mean spreads) * | -0.039 | -0.111 |
| BHL2 (mean parameters) * | -0.053 | -0.106 |
| BHL3 (negatives to be zero) * | 0.342 | 0.816 |
| CS1 (mean spreads) ^ | 0.027 | 0.31 |
| CS2 (mean parameters) ^ | -0.037 | 0.008 |
| CS3‡ (negatives to be zero) | 0.322 | 0.539 |
| AR‡ (negatives to be zero) | 0.269 | 0.697 |
| ROLL | 0.063 | 0.364 |
| Combination1 (BHL1+SHL2)/2 | -0.02 | 0.061 |
| Combination2 (SHL2+CS2)/2 | -0.019 | 0.292 |
| Combination3 (CS1+BHL1)/2 | -0.018 | -0.042 |
| Combination4 (BHL1+BHL2)/2 | -0.046 | -0.109 |

This table reports the average time-series and cross-sectional correlation of the estimators and the combinations.

Tick by tick quoted data are used to generate the monthly quoted spread.

Highest two correlation coefficients are made bold.

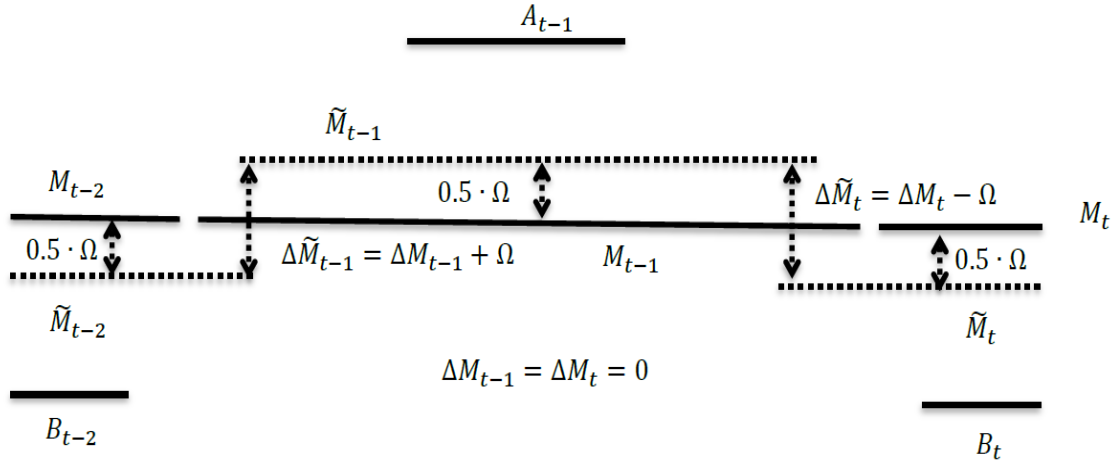


Figure 1: The Conjecture of the Spread

Figure 1 outlines the reasoning underpinning this proposition where for the purposes of economy we hold that the mid-price is fixed. The conjectural spread (\tilde{S}_t) is less than the true spread. This allows us to estimate the conjectural mid-price \tilde{M} ; this is represented by the dotted line in Figure 1, and the true mid price and transaction price are both represented by unbroken lines. Also in Figure 1, A and B denote observed ask and bid prices, whereas M is the unobserved true mid-price. In addition, Δ is taken to be the first-order difference operator and Ω denotes the conjectural error.

At any one point we can only observe one price, either the bid or the ask. In Figure 1, three periods are displayed. In the period labelled $t - 2$, the bid price is recorded and in period labelled $t - 1$, the ask price is observed. In period $t - 2$, the conjectural spread is lower than the true spread and the conjectural mid-price error is -0.5Ω , which is less than the true value. In period $t - 1$, the conjectural mid-price error is 0.5Ω , therefore this is greater than the true one. In the intervening period between $t - 2$ and $t - 1$, the direction of the trade shifts from sell to buy, and because of the conjectural error, we overestimate the mid-price return, formally we express this as:

$$\Delta \tilde{M}_{t-1} = \Delta M_{t-1} + \Omega = \Omega$$

In Figure 1, the hypothetical example shows that the variance of mid-price returns equates to zero because returns remain fixed. However the variance of conjectured mid-price returns is greater than zero. The reason for this is that in the case where the spread is underestimated, the conjectured mid-price fluctuates more than its true counterparts.

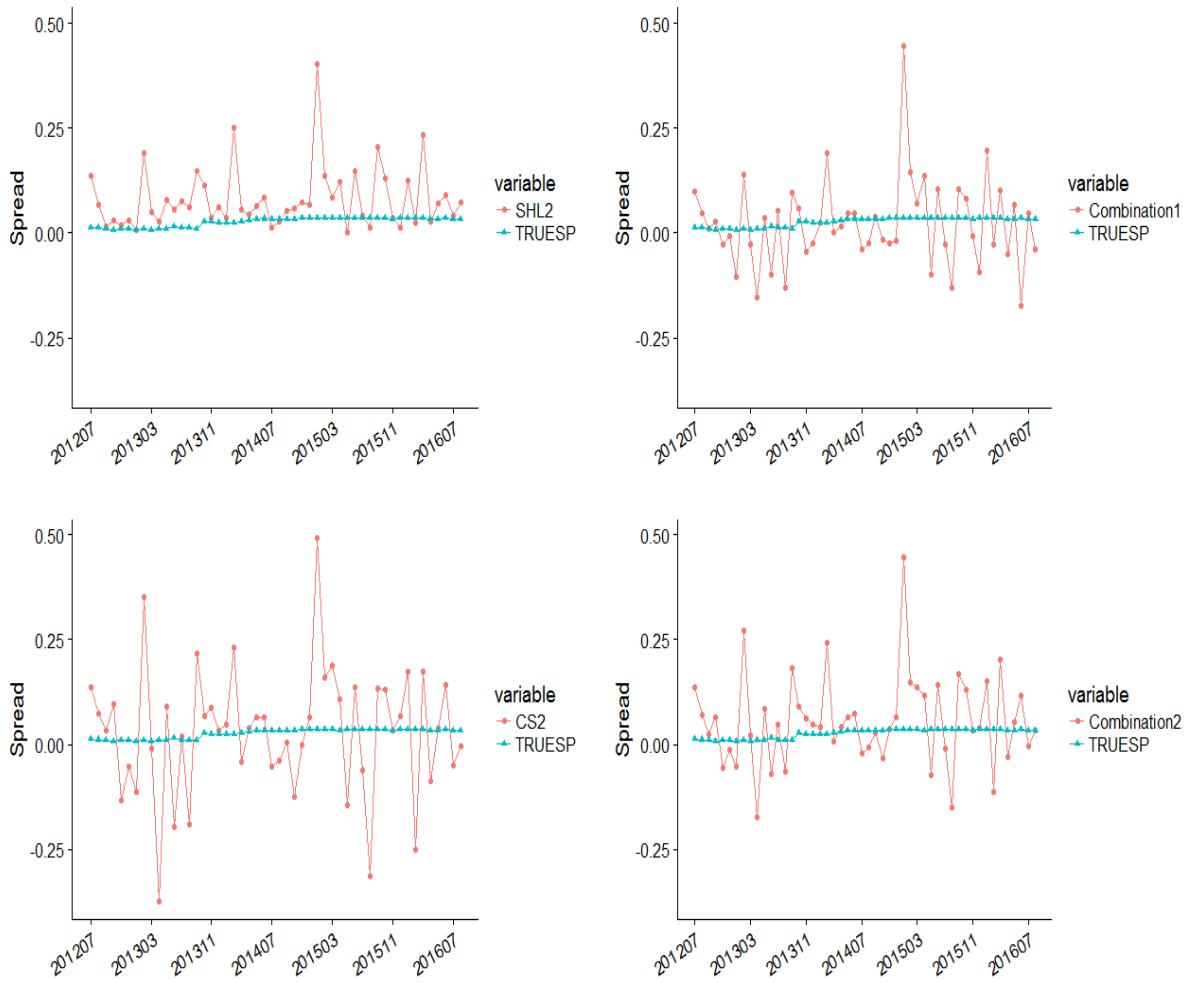


Figure 2: Monthly true and estimated spread, USD/JPY

The graphs above display estimates together with true values for a spread over a 50 month period from July 2012 to August 2016.

The currency pair chosen for the illustration without losing generality is USD/JPY. True spreads (TRUESP) data are monthly average closing spreads taken from DataStream.

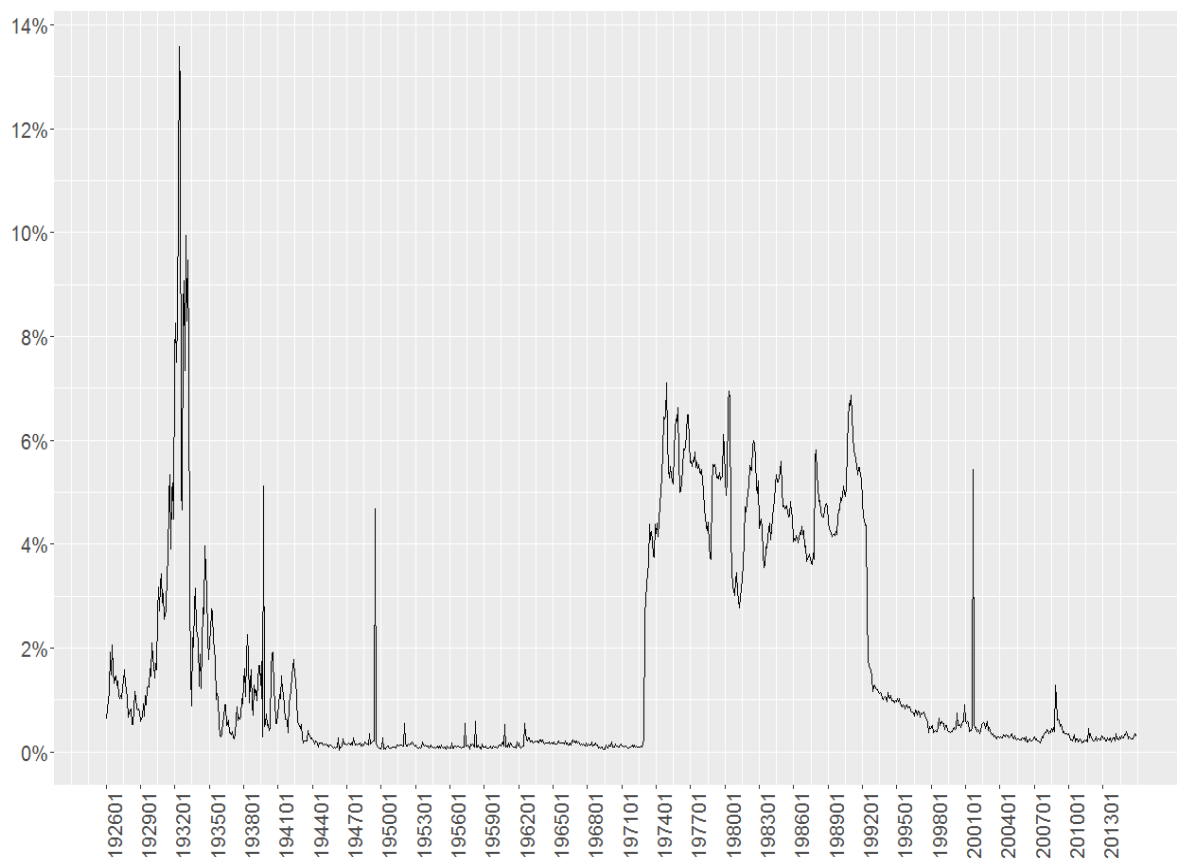


Figure 3: Monthly average estimated spread by SHL2 from 1926 to 2015, NYSE

Depicted here is SHL2 estimated bid-ask spreads for all stocks listed on the New York Stock Exchange on a monthly basis from January 1926 to December 2015. The figure plots the monthly equally weighted average spread of all stocks with each recording at least 16 daily spread observations within the month. All data is taken from CRSP.

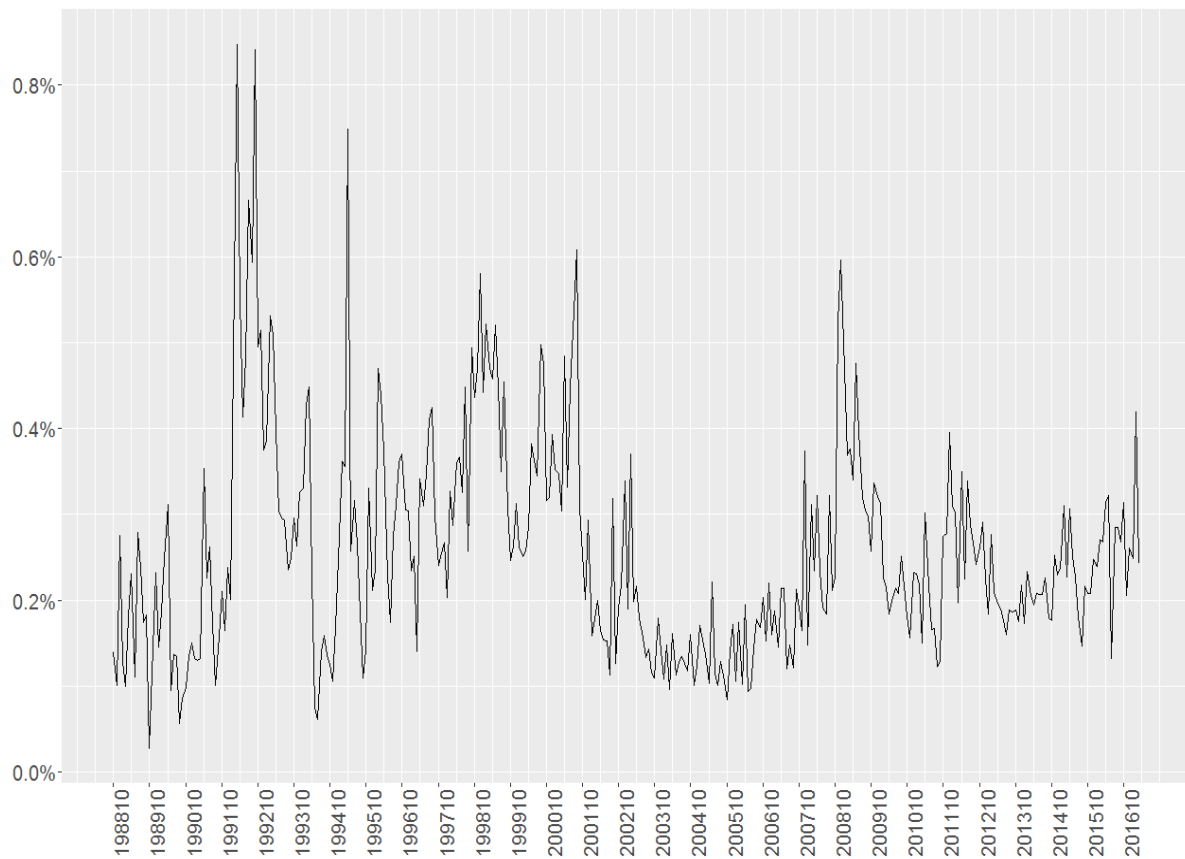


Figure 4: Monthly average estimated spread by SHL2 from 1990 to 2017,UK

Depicted here is SHL2 estimated bid-ask spreads for all stocks listed on the London Stock Exchange on a monthly basis from October 1988 to March 2017. The figure plots the monthly equally weighted average spread of all stocks with at least 16 daily spread observations within the month. All data is taken from Bloomberg.

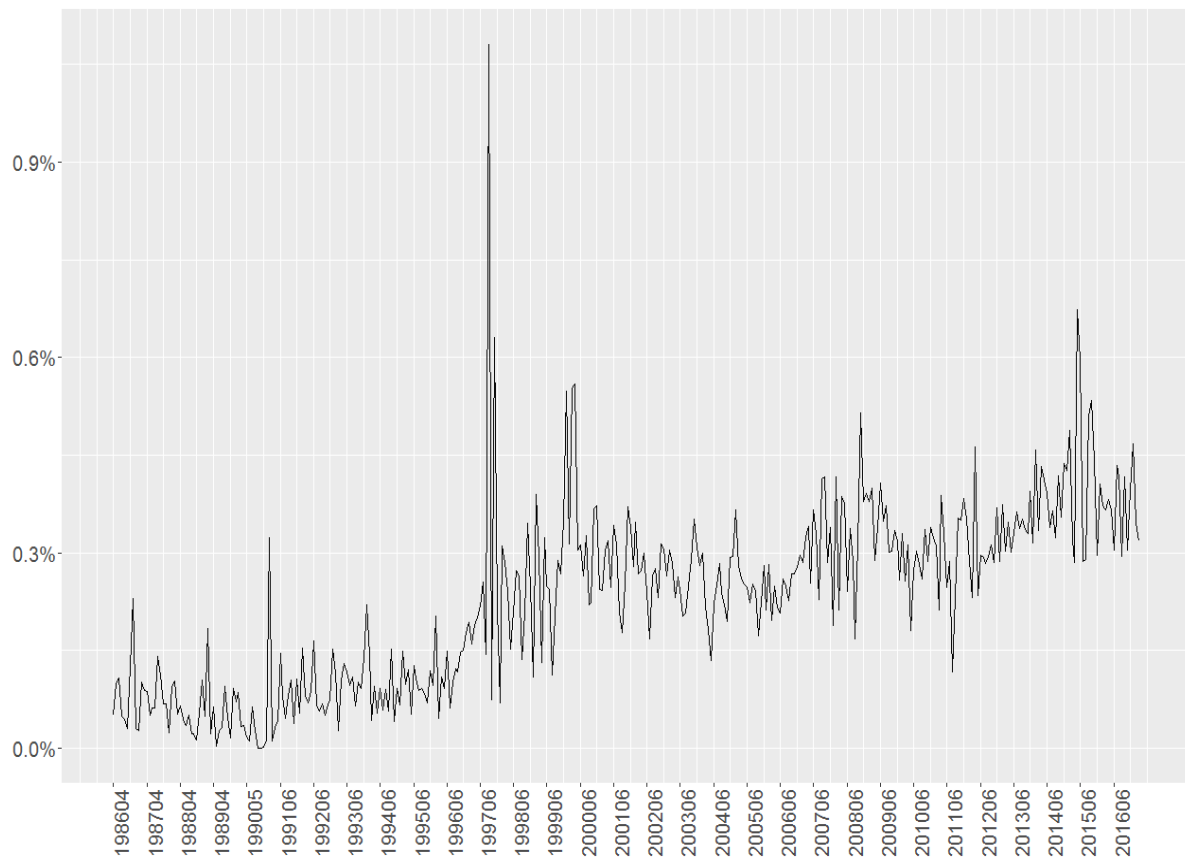


Figure 5: Monthly average estimated spread by SHL2 from 1986 to 2017, Hong Kong
 Depicted here is SHL2 estimated bid-ask spreads for all stocks listed on the Hong Kong Stock Exchange on a monthly basis from April 1986 to March 2017. The figure plots the monthly equally weighted average spread of all stocks with each recording at least 16 daily spread observations within the month. All data is taken from Bloomberg.

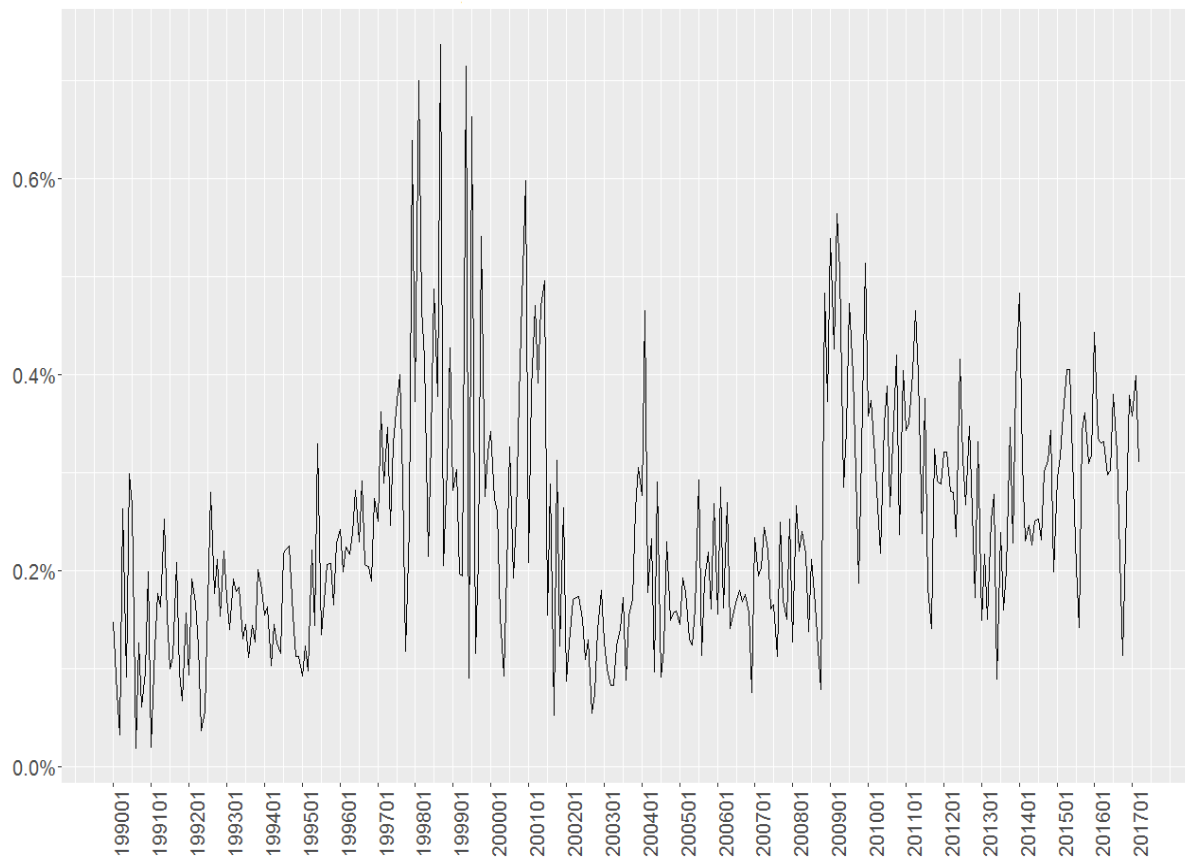


Figure 6: Monthly average estimated spread by SHL2 from 1990 to 2017, Thailand

Depicted here is SHL2 estimated bid-ask spreads for all stocks listed on the Stock Exchange of Thailand on a monthly basis from January 1990 to March 2017. The figure plots the monthly equally weighted average spread of all stocks with each recording at least 16 daily spread observations within the month. All data is taken from Bloomberg.